Convergence of perturbed Proximal Gradient algorithms

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Based on joint works with

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- Eric Moulines (Ecole Polytechnique, France)

↪ `On Perturbed Proximal-Gradient algorithms (JMLR, 2016)`

- Jean-François Aujol (IMB, Bordeaux, France)
- Charles Dossal (IMB, Bordeaux, France).

↪ `Acceleration for perturbed Proximal Gradient algorithms (work in progress)`

- Edouard Ollier (ENS Lyon, France)
- Adeline Samson (Univ. Grenoble Alpes, France).

↪ `Penalized inference in Mixed Models by Proximal Gradient methods (work in progress)`

Motivation : Pharmacokinetic (1/2)

- N patients.
- At time 0: dose D of a drug.
- For patient i, observations $\{Y_{ij}, 1 \leq j \leq J_i\}$: evolution of the concentration at times $t_{ij}, 1 \leq j \leq J_i$.

Model:

$$Y_{ij} = \mathcal{F}(t_{ij}, X_i) + \epsilon_{ij} \qquad \epsilon_{ij} \overset{i.i.d.}{\sim} \mathcal{N}(0, \sigma^2)$$

$X_i = Z_i \beta + d_i \in \mathbb{R}^L \qquad d_i \overset{i.i.d.}{\sim} \mathcal{N}_L(0, \Omega)$ and independent of ϵ•

Z_i known matrix s.t. each row of X_i has in intercept (fixed effect) and covariates

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Example of model $\mathcal{F}$: monocompartimental, oral administration

$$\mathcal{F}(t, [\ln \mathsf{Cl}, \ln \mathsf{V}, \ln \mathsf{A}]) = \mathcal{C}(\mathsf{Cl},\mathsf{V},\mathsf{A},\mathsf{D}) \left(\exp(-\frac{\mathsf{Cl}}{\mathsf{V}} t) - \exp(-\mathsf{A} t) \right)$$

For each patient i,

$$\begin{bmatrix} \ln Cl \\ \ln V \\ \ln A \end{bmatrix}_i = \begin{bmatrix} \beta_{0,Cl} \\ \beta_{0,V} \\ \beta_{0,A} \end{bmatrix} + \begin{bmatrix} \beta_{1,Cl} Z^i_{1,Cl} + \cdots + \beta_{K,Cl} Z^i_{K,Cl} \\ \text{idem, with covariates } Z^i_{k,V} \text{ and coefficients } \beta_{k,V} \\ \text{idem, with covariates } Z^i_{k,A} \text{ and coefficients } \beta_{k,A} \end{bmatrix} + \begin{bmatrix} d_{\mathsf{Cl},i} \\ d_{\mathsf{V},i} \\ d_{\mathsf{A},i} \end{bmatrix}$$

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Statistical analysis:

- estimation of $\theta = (\beta, \sigma^2, \Omega)$, under sparsity constraints on β
- selection of the covariates based on $\hat{\beta}$.

$\hookrightarrow$ Penalized Maximum Likelihood

Motivation : Pharmacokinetic (2/2)

Model:

$$Y_{ij} = f(t_{ij}, X_i) + \epsilon_{ij} \qquad \epsilon_{ij} \overset{i.i.d.}{\sim} \mathcal{N}(0, \sigma^2)$$

$$X_i = Z_i \beta + d_i \in \mathbb{R}^L \qquad d_i \overset{i.i.d.}{\sim} \mathcal{N}_L(0, \Omega) \text{ and independent of } \epsilon_\bullet$$

Z_i known matrix s.t. each row of X_i has in intercept (fixed effect) and covariates

Likelihoods:

- Likelihood: not explicit.
- Complete likelihood: the distribution of $\{Y_{ij}, X_i; 1 \leq i \leq N, 1 \leq j \leq J\}$ has an explicit expression.
- ML: here, the likelihood is not concave.

Outline

Penalized Maximum Likelihood inference with intractable Likelihood

- N observations : Y $= (Y_1, \cdots, Y_N)$
- A parametric statistical model $\theta \in \Theta \subseteq \mathbb{R}^d$ the dependance upon Y is omitted

$$\theta \mapsto L(\theta) \qquad \text{likelihood of the observations}$$

- A penalty term on the parameter θ: $\theta \mapsto g(\theta) \geq 0$ for sparsity constraints on θ. Usually, g non-smooth and convex.

Goal: Computation of

$$\theta \mapsto \operatorname{argmax}_{\theta \in \Theta} \left(\frac{1}{N} \log L(\theta) - g(\theta) \right)$$

when the likelihood L has no closed form expression, and can not be evaluated.

Example 1: Latent variable model

- The log-likelihood of the observations Y is of the form

$$\theta \mapsto \log L(\theta) \qquad L(\theta) = \int_{\mathsf{X}} p_\theta(x)\, \mu(\mathsf{d}x),$$

  where μ is a positive σ-finite measure on a set X.
- x collects the missing/latent data.

In these models,

- the complete likelihood $p_\theta(x)$ can be evaluated explicitly,
- the likelihood has no closed form expression.
- The exact integral could be replaced by a Monte Carlo approximation ; known to be inefficient.
  Numerical methods based on the a posteriori distribution of the missing data are preferred (see e.g. Expectation-Maximization approaches).

$\hookrightarrow$ What about the gradient of the (log)-likelihood ?

Gradient of the likelihood in a latent variable model

$$\log L(\theta) = \log \int_{\mathsf{X}} p_\theta(x)\, \mu(\mathsf{d}x)$$

Under regularity conditions, $\theta \mapsto \log L(\theta)$ is C^1 and

$$\nabla \log L(\theta) = \frac{\int_{\mathsf{X}} \partial_\theta p_\theta(x)\, \mu(\mathsf{d}x)}{\int_{\mathsf{X}} p_\theta(z)\, \mu(\mathsf{d}z)}$$
$$= \int_{\mathsf{X}} \partial_\theta \log p_\theta(x) \quad \underbrace{\frac{p_\theta(x)\, \mu(\mathsf{d}x)}{\int_{\mathsf{X}} p_\theta(z)\, \mu(\mathsf{d}z)}}_{\text{the a posteriori distribution}}$$

Gradient of the likelihood in a latent variable model

$$\log L(\theta) = \log \int_{\mathsf{X}} p_\theta(x)\, \mu(\mathrm{d}x)$$

Under regularity conditions, $\theta \mapsto \log L(\theta)$ is C^1 and

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$$= \int_{\mathsf{X}} \partial_\theta \log p_\theta(x) \quad \underbrace{\frac{p_\theta(x)\, \mu(\mathrm{d}x)}{\int_{\mathsf{X}} p_\theta(z)\, \mu(\mathrm{d}z)}}_{\text{the a posteriori distribution}}$$

The gradient of the log-likelihood

$$\nabla_\theta \left\{\log L(\theta)\right\} = \int_{\mathsf{X}} \partial_\theta \log p_\theta(x) \ \pi_\theta(\mathrm{d}x)$$

is an intractable expectation w.r.t. the conditional distribution of the latent variable given the observations Y*.*
For all (x, θ), $\partial_\theta \log p_\theta(x)$ *can be evaluated.*

Approximation of the gradient

$$\nabla_\theta \{\log L(\theta)\} = \int_{\mathsf{X}} \partial_\theta \log p_\theta(x) \; \pi_\theta(\mathrm{d}x)$$

1. Quadrature techniques: poor behavior w.r.t. the dimension of X
2. use i.i.d. samples from π_θ to define a Monte Carlo approximation: not possible, in general.
3. use m samples from a non stationary Markov chain $\{X_{j,\theta}, j \geq 0\}$ with unique stationary distribution π_θ, and define a Monte Carlo approximation. MCMC samplers provide such a chain.

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Stochastic approximation of the gradient

A biased approximation, since for MCMC samples $X_{j,\theta}$

$$\mathbb{E}\left[h(X_{j,\theta})\right] \neq \int h(x)\,\pi_\theta(\mathrm{d}x).$$

If the Markov chain is ergodic, the bias vanishes when $j \to \infty$.

Example 2: Discrete graphical model (Markov random field)

N independent observations of an undirected graph with p nodes.
Each node takes values in a finite alphabet X.

- N i.i.d. observations Y_i in X^p with distribution

$$y = (y_1, \cdots, y_p) \mapsto \pi_\theta(y) \stackrel{\text{def}}{=} \frac{1}{Z_\theta} \exp\left(\sum_{k=1}^{p} \theta_{kk} B(y_k, y_k) + \sum_{1 \leq j < k \leq p} \theta_{kj} B(y_k, y_j) \right)$$
$$= \frac{1}{Z_\theta} \exp\left(\langle \theta, \bar{B}(y) \rangle\right)$$

  where B is a symmetric function.
- θ is a symmetric $p \times p$ matrix.
- the normalizing constant (partition function) Z_θ can not be computed - sum over $|\mathsf{X}|^p$ terms.

Likelihood and its gradient in Markov random field

- Likelihood of the form (scalar product between matrices = Frobenius inner product)

$$\frac{1}{N}\log L(\theta) = \left\langle \theta, \frac{1}{N}\sum_{i=1}^{N} \bar{B}(Y_i) \right\rangle - \log Z_\theta$$

The likelihood is intractable.

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- Likelihood of the form (scalar product between matrices = Frobenius inner product)

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The likelihood is intractable.

- Gradient of the form

$$\nabla_\theta \left(\frac{1}{N} \log L(\theta) \right) = \frac{1}{N} \sum_{i=1}^{N} \bar{B}(Y_i) - \int_{\mathsf{X}^p} \bar{B}(y)\, \pi_\theta(y)\, \mu(\mathrm{d}y)$$

with

$$\pi_\theta(y) \stackrel{\text{def}}{=} \frac{1}{Z_\theta} \exp\left(\langle \theta, \bar{B}(y) \rangle\right).$$

The gradient of the (log)-likelihood is intractable.

Approximation of the gradient

$$\nabla_\theta \left(\frac{1}{N} \log L(\theta) \right) = \frac{1}{N} \sum_{i=1}^{N} \bar{B}(Y_i) - \int_{\mathsf{X}^p} \bar{B}(y)\, \pi_\theta(y)\, \mu(\mathrm{d}y).$$

The Gibbs measure

$$\pi_\theta(y) \stackrel{\text{def}}{=} \frac{1}{Z_\theta} \exp\left(\langle \theta, \bar{B}(y) \rangle\right)$$

is known up to the constant Z_θ.

Exact sampling from π_θ can be approximated by MCMC samplers (Gibbs-type samplers such as Swendsen-Wang, ...)

A biased approximation of the gradient is available.

To summarize,

Problem:

$$\mathrm{argmin}_{\theta \in \Theta} F(\theta) \qquad \text{with } F(\theta) = f(\theta) + g(\theta)$$

when

- $\theta \in \Theta \subseteq \mathbb{R}^d$
- the function g convex non-smooth nonnegative function (explicit)

To summarize,

Problem:

$$\mathrm{argmin}_{\theta \in \Theta} F(\theta) \qquad \text{with } F(\theta) = f(\theta) + g(\theta)$$

when

- $\theta \in \Theta \subseteq \mathbb{R}^d$
- the function g convex non-smooth nonnegative function (explicit)
- the function f is
  - · not necessarily convex,
  - · C^1 and ∇f is L-Lipschitz

$$\exists L > 0, \ \forall \theta, \theta' \qquad \|\nabla f(\theta) - \nabla f(\theta')\| \leq L \|\theta - \theta'\|.$$

  - · with an intractable gradient of the form

$$\nabla f(\theta) = \int H_\theta(x) \, \pi_\theta(\mathrm{d}x);$$

which can be approximated by **biased** Monte Carlo techniques.

Outline

The Proximal-Gradient algorithm (1/2)

$$\mathrm{argmin}_{\theta \in \Theta} F(\theta) \qquad \text{with } F(\theta) = \underbrace{f(\theta)}_{\text{smooth}} + \underbrace{g(\theta)}_{\text{non smooth}}$$

The Proximal Gradient algorithm

Given a stepsize sequence $\{\gamma_n, n \geq 0\}$, *iterative algorithm:*

$$\theta_{n+1} = \mathrm{Prox}_{\gamma_{n+1}, g}\left(\theta_n - \gamma_{n+1} \nabla f(\theta_n)\right)$$

where

$$\mathrm{Prox}_{\gamma, g}(\tau) \stackrel{\mathrm{def}}{=} \mathrm{argmin}_{\theta \in \Theta} \left(g(\theta) + \frac{1}{2\gamma} \|\theta - \tau\|^2 \right)$$

Proximal map: Moreau(1962)

Proximal Gradient algorithm: Beck-Teboulle(2010); Combettes-Pesquet(2011); Parikh-Boyd(2013)

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Proximal map: Moreau(1962)

Proximal Gradient algorithm: Beck-Teboulle(2010); Combettes-Pesquet(2011); Parikh-Boyd(2013)

- A generalization of the gradient algorithm to a composite objective function.
- A MM/Majorize-Minimize algorithm from a quadratic majorization of f (since Lipschitz gradient) which produces a sequence $\{\theta_n, n \geq 0\}$ such that

$$F(\theta_{n+1}) \leq F(\theta_n).$$

The proximal-gradient algorithm (2/2)

$$\mathrm{argmin}_{\theta\in\Theta} F(\theta) \qquad \text{with } F(\theta) = \underbrace{f(\theta)}_{\text{smooth}} + \underbrace{g(\theta)}_{\text{non smooth}}$$

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About the Prox-step:

- when $g = 0$: $\mathrm{Prox}(\tau) = \tau$
- when g is the $\{0, +\infty\}$-valued indicator fct of a closed convex set: the algorithm is the projected gradient.
- in some cases, Prox is explicit (e.g. elastic net penalty). Otherwise, numerical approximation:

$$\theta_{n+1} = \mathrm{Prox}_{\gamma_{n+1},g}\left(\theta_n - \gamma_{n+1}\nabla f(\theta_n)\right) + \epsilon_{n+1}$$

in this talk, $\epsilon_{n+1} = 0$

The perturbed proximal-gradient algorithm

The Perturbed Proximal Gradient algorithm

Given a stepsize sequence $\{\gamma_n, n \geq 0\}$*, iterative algorithm:*

$$\theta_{n+1} = \mathrm{Prox}_{\gamma_{n+1},g}\left(\theta_n - \gamma_{n+1}\mathbf{H_{n+1}}\right)$$

where H_{n+1} *is an approximation of* $\nabla f(\theta_n)$*.*

Monte Carlo-Proximal Gradient algorithm

In the case:

$$\nabla f(\theta) = \int H_\theta(x)\, \pi_\theta(x) \mu(\mathrm{d}x),$$

The MC-Proximal Gradient algorithm

Choose a stepsize sequence $\{\gamma_n, n \geq 0\}$ *and a batch size sequence* $\{m_n, n \geq 0\}$.

Given the current value θ_n,

1. *Sample a Markov chain* $\{X_{j,n}, j \geq 0\}$ *from a MCMC sampler with kernel* $P_{\theta_n}(x, \mathrm{d}x')$, *and unique invariant distribution* $\pi_{\theta_n}\, \mathrm{d}\mu$.
2. *Set*

$$H_{n+1} = \frac{1}{m_{n+1}} \sum_{j=1}^{m_{n+1}} H_{\theta_n}(X_{j,n}).$$

3. *Update the value of the parameter*

$$\theta_{n+1} = \mathrm{Prox}_{\gamma_{n+1},g}\left(\theta_n - \gamma_{n+1} H_{n+1}\right)$$

Stochastic Approximation-Proximal Gradient algorithm

In the case (ex. latent variable models with exponential complete likelihood;log-linear Markov random field)

$$\nabla f(\theta) = \int H_\theta(x)\, \pi_\theta(x) \mu(\mathrm{d}x), \qquad H_\theta(x) = \Phi(\theta) + \Psi(\theta) S(x)$$

which implies

$$\nabla f(\theta) = \Phi(\theta) + \Psi(\theta) \left(\int S(x)\, \pi_\theta(x) \mu(\mathrm{d}x) \right),$$

The SA-Proximal Gradient algorithm

Choose two stepsize sequences $\{\gamma_n, \delta_n, n \geq 0\}$ and a batch size sequence $\{m_n, n \geq 0\}$

Given the current value θ_n,

1. *Sample a Markov chain $\{X_{j,n}, j \geq 0\}$ from a MCMC sampler with kernel $P_{\theta_n}(x, \mathrm{d}x')$, and unique invariant distribution $\pi_{\theta_n} \,\mathrm{d}\mu$.*
2. *Set $H_{n+1} = \Phi(\theta_n) + \Psi(\theta_n) S_{n+1}$ with*

$$S_{n+1} = (1 - \delta_{n+1})\, S_n + \delta_{n+1} \frac{1}{m_{n+1}} \sum_{j=1}^{m_{n+1}} S(X_{j,n}).$$

3. *Update the value of the parameter*

(*) Penalized Expectation-Maximization (EM) vs Proximal-Gradient

- EM Dempster et al. (1977) is a Majorize-Minimize algorithm for the computation of the ML estimate in latent variable models.
- Penalized (Stochastic) EM algorithms

$$
\begin{aligned}
\tau_{n+1} &= \mathrm{argmax}_\theta \int \log p_\theta(x)\ \pi_\theta(x)\, \mathsf{d}\mu(x) - g(\theta) \\
&= \mathrm{argmax}_\theta \left\{ A(\theta) + \langle B(\theta), S_{n+1} \rangle - g(\theta) \right\}
\end{aligned}
$$

with

$$S_{n+1} = \int S(x)\ \pi_{\tau_n}(x)\, \mathsf{d}\mu(x) \qquad \text{EM}$$

$$S_{n+1} = \frac{1}{m_{n+1}} \sum_{j=1}^{m_{n+1}} S(X_{j,n}) \qquad \text{Monte Carlo EM}$$ Wei and Tanner (1990)

$$S_{n+1} = (1-\delta_{n+1}) S_n + \frac{\delta_{n+1}}{m_{n+1}} \sum_{j=1}^{m_{n+1}} S(X_{j,n}) \qquad \text{Stoch. Approx. EM}$$ Delyon et al. (1999)

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or choose τ_{n+1} s.t.

$$A(\tau_{n+1}) + \langle B(\tau_{n+1}), S_{n+1} \rangle - g(\tau_{n+1}) \geq A(\tau_n) + \langle B(\tau_n), S_{n+1} \rangle - g(\tau_n)$$

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$$S_{n+1} = \int S(x)\ \pi_{\tau_n}(x)\, \mathsf{d}\mu(x) \qquad \text{EM}$$

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with

- MC-Prox Gdt and SA-Prox Gdt are Penalized Stochastic Generalized EM algorithms.

Numerical illustration (1/3): pharmacokinetic

For the implementation of the algorithm

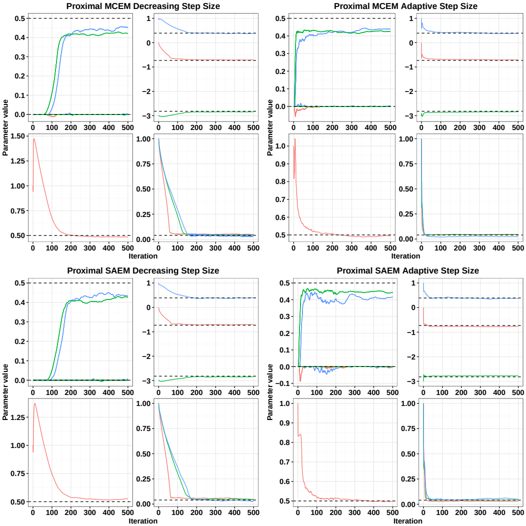
- Penalty term: $g(\theta) = \lambda \|\beta\|_1$. How to choose λ ?

  $\hookrightarrow \lambda = \mathrm{argmin}_{\lambda_1, \cdots, \lambda_L}$ E-BIC$(\hat{\beta}_\lambda)$

- Stepsize sequences: constant or vanishing stepsize sequence $\{\gamma_n, n \geq 0\}$? (and δ_n for the SA-Prox Gdt algorithm)

- Monte Carlo approximation: fixed or increasing batch size ?

Numerical illustration (2/3): pharmacokinetic



Numerical illustration (3/3): pharmacokinetic

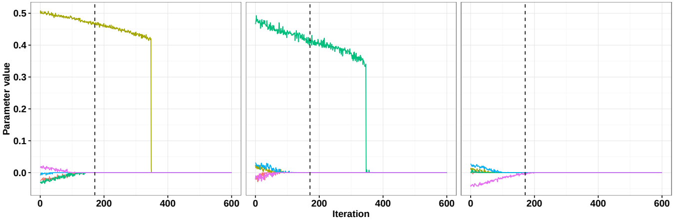


Figure: Regularization path of the covariate parameters for the clearance (left), absorption constant (middle) and volume of distribution (right) parameters. Black dashed line corresponds to the λ value selected by $EBIC$. Each colored curve corresponds to a covariate.

Outline

The assumptions

$$\mathrm{argmin}_{\theta \in \Theta} F(\theta) \qquad \text{with } F(\theta) = f(\theta) + g(\theta)$$

where

- the function g: $\mathbb{R}^d \to [0, \infty]$ is convex, non smooth, not identically equal to $+\infty$, and lower semi-continuous

- the function f: $\mathbb{R}^d \to \mathbb{R}$ is a smooth **convex** function

  i.e. f is continuously differentiable and there exists $L > 0$ such that

$$\|\nabla f(\theta) - \nabla f(\theta')\| \leq L \, \|\theta - \theta'\| \qquad \forall \theta, \theta' \in \mathbb{R}^d$$

- $\Theta \subseteq \mathbb{R}^d$ is the domain of g: $\Theta = \{\theta \in \mathbb{R}^d : g(\theta) < \infty\}$.
- The set $\mathrm{argmin}_\Theta F$ is a non-empty subset of Θ.

Existing results in the literature

There exist results under (some of) the assumptions

$$\mathbb{E}\left[H_{n+1}|\mathcal{F}_n\right] = \nabla f(\theta_n), \qquad \inf_n \gamma_n > 0, \qquad \sum_n \|H_{n+1} - \nabla f(\theta_n)\| < \infty,$$

i.e. results for

- unbiased sampling. Almost no conditions for the biased sampling, such as the MCMC one.
- non vanishing stepsize sequence $\{\gamma_n, n \geq 0\}$.
- increasing batch size: when H_{n+1} is a Monte Carlo sum i.e.

$$H_{n+1} = \frac{1}{m_{n+1}} \sum_{j=1}^{m_{n+1}} H_{\theta_n}(X_{j,n}),$$

the assumptions imply that $\lim_n m_n = +\infty$ at some rate.

Combettes (2001) Elsevier Science.
Combettes-Wajs (2005) Multiscale Modeling and Simulation.
Combettes-Pesquet (2015, 2016) SIAM J. Optim, arXiv
Lin-Rosasco-Villa-Zhou (2015) arXiv
Rosasco-Villa-Vu (2014,2015) arXiv
Schmidt-Leroux-Bach (2011) NIPS

Convergence of the perturbed proximal gradient algorithm (1/3)

$$\theta_{n+1} = \mathrm{Prox}_{\gamma_{n+1},g}\left(\theta_n - \gamma_{n+1}\ H_{n+1}\right) \qquad \text{with } H_{n+1} \approx \nabla f(\theta_n)$$

Set: $\mathcal{L} = \mathrm{argmin}_{\Theta}(f+g)$ $\qquad \eta_{n+1} = H_{n+1} - \nabla f(\theta_n)$

Theorem (Atchadé, F., Moulines (2015))

Assume

- g *convex, lower semi-continuous;* f *convex,* C^1 *and its gradient is Lipschitz with constant* L*;* $\mathcal{L}$ *is non empty.*
- $\sum_n \gamma_n = +\infty$ *and* $\gamma_n \in (0, 1/L]$.
- *Convergence of the series*

$$\sum_n \gamma_{n+1}^2 \|\eta_{n+1}\|^2, \qquad \sum_n \gamma_{n+1}\eta_{n+1}, \qquad \sum_n \gamma_{n+1} \left\langle \mathsf{T}_n, \eta_{n+1} \right\rangle$$

where $\mathsf{T}_n = \mathrm{Prox}_{\gamma_{n+1},g}(\theta_n - \gamma_{n+1}\nabla f(\theta_n))$.

Then there exists $\theta_\star \in \mathcal{L}$ *such that* $\lim_n \theta_n = \theta_\star$.

Convergence of the perturbed proximal gradient algorithm (2/3)

This convergence result

- for the convex case: f and g are convex.

Convergence of the perturbed proximal gradient algorithm (2/3)

This convergence result

- for the convex case: f and g are convex.
- is a deterministic result.
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1. Applications in Computational Statistics

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1. Applications in Computational Statistics
2. Applications in learning - "finite sum context" :

(objective) $$\mathrm{argmin}_\theta \left(\frac{1}{N} \sum_{i=1}^{N} f_i(\theta) + g(\theta) \right)$$

(Approx. Gdt) $$H_{n+1} = \frac{1}{|I_{n+1}|} \sum_{i \in I_{n+1}} \nabla f_i(\theta_n)$$

(X_i's) the indices $i \in I_{n+1}$

Proof / Convergence of the perturbed proximal gradient algorithm (3/3)

Its proof relies on

1. a deterministic Lyapunov inequality

$$\|\theta_{n+1}-\theta_\star\|^2 \leq \|\theta_n-\theta_\star\|^2 - \underbrace{2\gamma_{n+1}\left(F(\theta_{n+1}) - \min F\right)}_{\text{non-negative}} \underbrace{-2\gamma_{n+1}\left\langle \mathsf{T}_n - \theta_\star, \eta_{n+1}\right\rangle + 2\gamma_{n+1}^2\|\eta_{n+1}\|^2}_{\text{signed noise}}$$

2. (an extension of) the Robbins-Siegmund lemma

Let $\{v_n, n \geq 0\}$ and $\{\chi_n, n \geq 0\}$ be non-negative sequences and $\{\xi_n, n \geq 0\}$ be such that $\sum_n \xi_n$ exists. If for any $n \geq 0$,

$$v_{n+1} \leq v_n - \chi_{n+1} + \xi_{n+1}$$

then $\sum_n \chi_n < \infty$ and $\lim_n v_n$ exists.

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Note: deterministic lemma, signed noise.

Convergence: when H_{n+1} is a Monte-Carlo approximation (1/3)

In the case

$$\nabla f(\theta_n) \approx H_{n+1} = \frac{1}{m_{n+1}} \sum_{j=1}^{m_{n+1}} H_{\theta_n}(X_{j,n}),$$

$$X_{j+1,n}|\text{past} \sim P_{\theta_n}(X_{j,n}, \cdot) \qquad \pi_\theta P_\theta = \pi_\theta;$$

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let us check the condition "$\sum_n \gamma_n \eta_n < \infty$ w.p.1" under the condition $\sum_n \gamma_n = +\infty$:

$$\begin{aligned}\sum_n \gamma_{n+1} \eta_{n+1} &= \sum_n \gamma_{n+1} \left(H_{n+1} - \nabla f(\theta_n)\right) \\ &= \sum_n \gamma_{n+1} \left\{H_{n+1} - \mathbb{E}\left[H_{n+1} | \mathcal{F}_n\right]\right\} + \sum_n \gamma_{n+1} \underbrace{\left\{\mathbb{E}\left[H_{n+1} | \mathcal{F}_n\right] - \nabla f(\theta_n)\right\}}_{\substack{\text{if unbiased MC: null} \\ \text{if biased MC: } O(1/m_n)}}\end{aligned}$$

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$$= \sum_n \gamma_{n+1}\left\{H_{n+1} - \mathbb{E}\left[H_{n+1}|\mathcal{F}_n\right]\right\} + \sum_n \gamma_{n+1} \underbrace{\left\{\mathbb{E}\left[H_{n+1}|\mathcal{F}_n\right] - \nabla f(\theta_n)\right\}}_{\substack{\text{if unbiased MC: null} \\ \text{if biased MC: } O(1/m_n)}}$$

The most technical case: the biased case with constant batch size $m_n = m$

Solution $\widehat{H}_\theta$ to the Poisson equation: $H_\theta - \pi_\theta H_\theta = \widehat{H}_\theta - P_\theta \widehat{H}_\theta$

$H_{n+1} - \nabla f(\theta_n) =$ martingale increment + remainder

Regularity in θ of $\theta \mapsto \widehat{H}_\theta$ and $\theta \mapsto P_\theta \widehat{H}_\theta$.

Convergence: when H_{n+1} is a Monte-Carlo approximation (2/3)

Increasing batch size: $\lim_n m_n = +\infty$

Conditions on the step sizes and batch sizes

$$\sum_n \gamma_n = +\infty, \qquad \sum_n \frac{\gamma_n^2}{m_n} < \infty; \qquad \sum_n \frac{\gamma_n}{m_n} < \infty \text{ (biased case)}$$

Conditions on the Markov kernels: *There exist* $\lambda \in (0,1)$, $b < \infty$, $p \geq 2$ *and a measurable function* $W : \mathsf{X} \to [1, +\infty)$ *such that*

$$\sup_{\theta \in \Theta} |H_\theta|_W < \infty, \qquad \sup_{\theta \in \Theta} P_\theta W^p \leq \lambda W^p + b.$$

In addition, for any $\ell \in (0, p]$, *there exist* $C < \infty$ *and* $\rho \in (0,1)$ *such that for any* $x \in \mathsf{X}$,

$$\sup_{\theta \in \Theta} \|P_\theta^n(x, \cdot) - \pi_\theta\|_{W^\ell} \leq C \rho^n W^\ell(x). \tag{1}$$

Condition on Θ: Θ *is bounded.*

Convergence: when H_{n+1} is a Monte-Carlo approximation (3/3)

Fixed batch size: $m_n = m$

Condition on the step size:

$$\sum_n \gamma_n = +\infty \qquad \sum_n \gamma_n^2 < \infty \qquad \sum_n |\gamma_{n+1} - \gamma_n| < \infty$$

Condition on the Markov chain: *same as in the case "increasing batch size" and there exists a constant C such that for any $\theta, \theta' \in \Theta$*

$$|H_\theta - H_{\theta'}|_W + \sup_x \frac{\|P_\theta(x,\cdot) - P_{\theta'}(x,\cdot)\|_W}{W(x)} + \|\pi_\theta - \pi_{\theta'}\|_W \leq C \, \|\theta - \theta'\|.$$

Condition on the Prox:

$$\sup_{\gamma \in (0,1/L]} \sup_{\theta \in \Theta} \gamma^{-1} \, \|\mathrm{Prox}_{\gamma,g}(\theta) - \theta\| < \infty.$$

Condition on Θ: Θ is bounded.

Rates of convergence (1/3) : the problem

For non negative weights a_k, find an upper bound of

$$\sum_{k=1}^{n} \frac{a_k}{\sum_{\ell=1}^{n} a_\ell} F(\theta_k) - \min F$$

It provides

- an upper bound for the cumulative regret $(a_k = 1)$
- an upper bound for an averaging strategy when F is convex since

$$F\left(\sum_{k=1}^{n} \frac{a_k}{\sum_{\ell=1}^{n} a_\ell} \theta_k\right) - \min F \leq \sum_{k=1}^{n} \frac{a_k}{\sum_{\ell=1}^{n} a_\ell} \ F(\theta_k) - \min F.$$

Rates of convergence (2/3): a deterministic control

Theorem (Atchadé, F., Moulines (2016))

For any $\theta_\star \in \mathrm{argmin}_\Theta F$,

$$\sum_{k=1}^{n} \frac{a_k}{A_n} F(\theta_k) - \min F \leq \frac{a_0}{2\gamma_0 A_n} \|\theta_0 - \theta_\star\|^2 + \frac{1}{2A_n} \sum_{k=1}^{n} \left(\frac{a_k}{\gamma_k} - \frac{a_{k-1}}{\gamma_{k-1}} \right) \|\theta_{k-1} - \theta_\star\|^2 + \frac{1}{A_n} \sum_{k=1}^{n} a_k \gamma_k \|\eta_k\|^2 - \frac{1}{A_n} \sum_{k=1}^{n} a_k \left\langle \mathsf{T}_{k-1} - \theta_\star, \eta_k \right\rangle$$

where

$$A_n = \sum_{\ell=1}^{n} a_\ell, \qquad \eta_k = H_k - \nabla f(\theta_{k-1}), \qquad \mathsf{T}_k = \mathrm{Prox}_{\gamma_k, g}(\theta_{k-1} - \gamma_k \nabla f(\theta_{k-1})).$$

Rates (3/3): when H_{n+1} is a Monte Carlo approximation, bound in L^q

$$\left\| F\left(\frac{1}{n}\sum_{k=1}^{n}\theta_k\right) - \min F \right\|_{L^q} \leq \left\| \frac{1}{n}\sum_{k=1}^{n} F(\theta_k) - \min F \right\|_{L^q} \leq u_n$$

$u_n = O(1/\sqrt{n})$

with fixed size of the batch and (slowly) decaying stepsize

$$\gamma_n = \frac{\gamma_\star}{n^a}, a \in [1/2, 1] \qquad m_n = m_\star.$$

With averaging: optimal rate, even with slowly decaying stepsize $\gamma_n \sim 1/\sqrt{n}$.

$u_n = O(\ln n/n)$

with increasing batch size and constant stepsize

$$\gamma_n = \gamma_\star \qquad m_n \propto n.$$

Rate with $O(n^2)$ *Monte Carlo samples !*

Acceleration (1)

Let $\{t_n, n \geq 0\}$ be a positive sequence s.t.

$$\gamma_{n+1} t_n (t_n - 1) \leq \gamma_n t_{n-1}^2$$

Nesterov acceleration of the Proximal Gradient algorithm

$$\theta_{n+1} = \mathrm{Prox}_{\gamma_{n+1},g} \left(\tau_n - \gamma_{n+1} \nabla f(\tau_n) \right)$$
$$\tau_{n+1} = \theta_{n+1} + \frac{t_n - 1}{t_{n+1}} \left(\theta_{n+1} - \theta_n \right)$$

Nesterov(2004), Tseng(2008), Beck-Teboulle(2009)

Zhu-Orecchia (2015); Attouch-Peypouquet(2015); Bubeck-Lee-Singh(2015); Su-Boyd-Candes(2015)

(deterministic) Proximal-gradient $\quad F(\theta_n) - \min F = O\left(\frac{1}{n}\right)$

(deterministic) Accelerated Proximal-gradient $\quad F(\theta_n) - \min F = O\left(\frac{1}{n^2}\right)$

Acceleration (2) Aujol-Dossal-F.-Moulines, work in progress

Perturbed Nesterov acceleration: some convergence results

Choose γ_n, m_n, t_n *s.t.*

$$\gamma_n \in (0, 1/L] \,, \qquad \lim_n \gamma_n t_n^2 = +\infty, \qquad \sum_n \gamma_n t_n (1 + \gamma_n t_n) \frac{1}{m_n} < \infty$$

Then there exists $\theta_\star \in \operatorname{argmin}_\Theta F$ *s.t* $\lim_n \theta_n = \theta_\star$.
In addition

$$F(\theta_{n+1}) - \min F = O\left(\frac{1}{\gamma_{n+1} t_n^2}\right)$$

Schmidt-Le Roux-Bach (2011); Dossal-Chambolle(2014); Aujol-Dossal(2015)

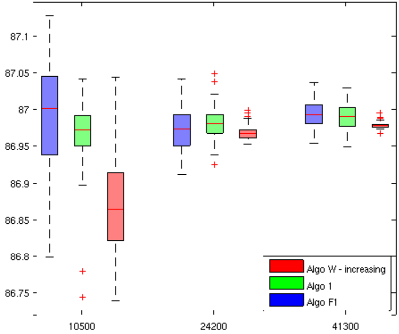
γ_n	m_n	t_n	rate	NbrMC
γ	n^3	n	n^{-2}	n^4
$\gamma/\sqrt{n}$	n^2	n	$n^{-3/2}$	n^3

Table: Control of $F(\theta_n) - \min F$

Outline

Conclusion (1/2): acceleration ?

- with or without the acceleration: complexity $O(1/\sqrt{n})$.
- acceleration: longer Markov chains, few iterations.



Conclusion (2/2): weaken the assumptions

- $\theta \in \mathbb{R}^d \to \theta$ in a Hilbert space
- Θ bounded $\to$ no boundedness condition on Θ
- f convex $\to$ f non convex