

Estimation of cosmological parameters using adaptive importance sampling

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Exploration du modèle cosmologique par fusion statistique de grands relevés hétérogènes

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Introduction

Objectives of the ANR project:

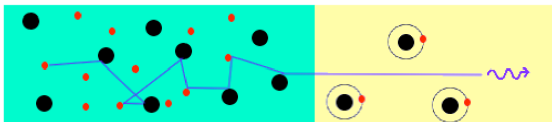
Combine three deep surveys of the universe to **set new constraints on the evolution scenario** of galaxies and large scale structures, **and the fundamental cosmological parameters**.

1. What is the "evolution scenario" ?
2. Examples of "cosmological parameters"
3. Example of data: Cosmic Microwave Background (CMB)

Evolution scenario of the Universe

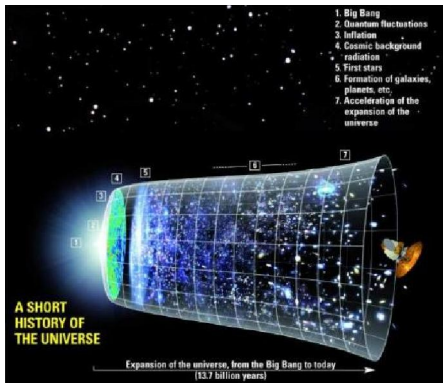
- The **cosmology** is the astrophysical study of the history and structure of the universe.

During the XXth century, a new theory for the *expansion* of the universe: **Cosmological Standard Model**. Therefore, today, cosmology also includes the study of the constituent dynamics of the universe.



- Expansion of the universe
 - from an extremely dense and hot state, a plasma of protons, electrons, photons, closely interacting with each other and in thermal equilibrium
 - to the vast and much cooler cosmos we currently have.
- Cooling down: thermal agitation can not prevent atoms to be formed.
- End of opaque universe: after recombination, matter and radiation

Open questions:



- will the universe expand for ever, or will it collapse?
- what is the shape of the universe?
- Is the expansion of the universe accelerating rather than decelerating?
- Is the universe dominated by dark matter and what is its concentration?

Cosmological parameters (I)

Description of the expansion by a scaling factor $a(t)$ of the space coordinates.

Example: Friedmann equations for the expansion model $a(t)$:

$$\left(\frac{a'(t)}{a(t)}\right)^2 = -\frac{K}{a^2(t)} + \frac{8\pi G}{3}\rho(t) + \frac{\Lambda}{3}$$

Solutions as a function of the spatial curvature K and the cosmological constant Λ

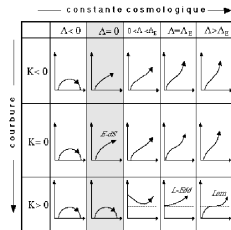


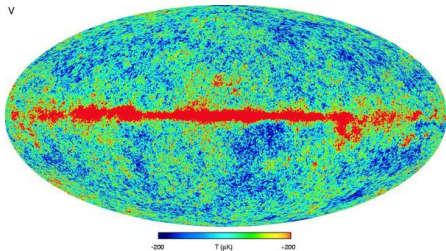
FIG.: Big Bang / Big crunch

Cosmological parameters (II)

- **Density of barionic matter Ω_b** This is ordinary matter composed of protons, neutrons, and electrons. It comprises gas, dust, stars, planets, ...
- **Density of cold dark matter Ω_c** It comprises the dark matter halos that surround galaxies and galaxy clusters, and aids in the formation of structure in the universe.
- **Density of dark energy Ω_Λ** (cosmological constant Λ) Through observations of distant supernovae, it was discovered that the expansion of the universe appears to be getting faster with time. Whatever the source of this phenomenon turns out to be, cosmologists refer to it generically as dark energy.
- **Hubble constant H_0**
- **Shape of the Universe K** (spatial curvature).

Data set 1: Cosmic Microwave Background

CMB is the radiation left over from an early stage in the development of the universe: after the *recombination epoch* when neutral atoms formed from protons and electrons, followed by the *photon decoupling* when photons started to travel freely through space.



Example of survey: WMAP for the **Cosmic Microwave Background (CMB)** radiations = temperature variations are related to fluctuations in the density of matter in the early universe and thus carry out information about the initial conditions for the formation of cosmic structures such as galaxies, clusters and voids for example.

From a CMB map to the cosmological parameters

- **Step 1: map making process**, from scanning the sky to producing spherical CMB maps
 - must exploit multi-scan, deal with asymmetric instrumental beams, ...
 - source separation, to remove "foreground emissions" (from galactic and extra-galactic origins)
- **Step 2: a likelihood function** to express the probability of a given CMB map given an angular power spectrum \mathcal{C} .
- **Step 3: a cosmological model** predicting the dependence of the angular power spectrum on the cosmological parameters; $\theta \mapsto \mathcal{C}(\theta)$.
↔ software packages (ex. CAMB, CMBfast).

(step 2) Likelihood function for a CMB map

The CMB map is the realization of a random process X on the unit sphere, X assumed to be stationary:

$$\text{Cov}(X(\xi), X(\xi')) = \rho(\xi^\dagger \xi') \quad \rho: \text{angular correlation function}$$

ρ is related to the angular power spectrum $\{C_\ell\}_\ell$

$$\rho(z) = \sum_{\ell \geq 0} C_\ell \frac{2\ell + 1}{4\pi} P_\ell(z) \quad P_\ell, \ell\text{-th Legendre polynomial}$$

How to estimate $\{C_\ell, \ell \geq 0\}$?

- Multipole decomposition:

$$X(\xi) = \sum_{\ell \geq 0} X^{(\ell)}(\xi) \quad \ell: \text{angular frequency}$$

- Define the empirical angular spectrum $\hat{C}_\ell = \frac{1}{2\ell+1} \|X^\ell\|^2$
 X^ℓ obtained by spherical convolution of X with P_ℓ .

The likelihood of the signal given the cosmological parameters i.e.

- signal \rightarrow summarized by the **empirical angular spectrum** $\{\hat{C}_\ell, \ell \geq 0\}$ - a kind of “sufficient statistics”
- cosmological parameters θ yielding to the theoretical **angular spectrum** $\{C_\ell(\theta), \ell \geq 0\}$ - by software packages

is given by

$$-2 \log p(\text{CMBmap}|\theta) = \sum_{\ell \geq 0} (2\ell + 1) \left(\frac{\hat{C}_\ell}{C_\ell(\theta)} + \log C_\ell(\theta) \right) + \text{Cst}$$

Reference: J.F. Cardoso, *Precision Cosmology with the Cosmic Microwave Background*, IEEE Signal Processing Magazine, 2010

Combining data sets

- Observational data from
 - the CMB Cosmic Microwave Background → five-year WMAP data.
 - the observation of weak gravitational shear → CFHTLS-Wide third release.
- explained by some cosmologic parameters

TABLE II: Parameters for the cosmology likelihood. C=CMB, S=SNiA, L=lensing.

Symbol	Description	Minimum	Maximum	Experiment
Ω_b	Baryon density	0.01	0.1	C L
Ω_m	Total matter density	0.01	1.2	C S L
w	Dark-energy eq. of state	-3.0	0.5	C S L
n_s	Primordial spectral index	0.7	1.4	C L
Δ_R^2	Normalization (large scales)			C
σ_8	Normalization (small scales) ^a			C L
h	Hubble constant			C L
τ	Optical depth			C
M	Absolute SNiA magnitude			S
α	Colour response			S
β	Stretch response			S
a				L
b	galaxy z -distribution fit			L
c				L

A challenging (a posteriori) density exploration

This yields:

- a likelihood of the data given the parameters: some of them computed from publicly available codes ex. WMAP5 code for CMB data
- combined with a priori knowledge: uniform prior on hypercubes.

Therefore, statistical inference consists in the exploration of the **a posteriori density** of the parameters, **a challenging task** due to

- potentially high dimensional parameter space (not really considered here: sampling in \mathbb{R}^d , $d \sim 10$ to 15)
- immensely slow computation of likelihoods,
- non-linear dependence and degeneracies between parameters introduced by physical constraints or theoretical assumptions.

II. Monte Carlo algorithms for the exploration of a (a posteriori) density π

Monte Carlo algorithms

- **(naive) Monte Carlo** methods: i.i.d. samples under π . Here, NO: π is only known through a "numerical box"
- **Importance Sampling** methods: i.i.d. samples $\{X_k, k \geq 0\}$ under a proposal distribution q and

$$\sum_{k=1}^n \frac{\omega_k}{\sum_{j=1}^n \omega_j} \mathbb{1}_{\Delta}(X_k) \approx \mathbb{P}_{\pi}(X \in \Delta) \quad \text{with} \quad \omega_k = \frac{\pi(X_k)}{q(X_k)}$$

- **Markov chain Monte Carlo** methods: a Markov chain with stationary distribution π

$$\frac{1}{n} \sum_{k=1}^n \mathbb{1}_{\Delta}(X_k) \approx \mathbb{P}_{\pi}(X \in \Delta)$$

Importance sampling or MCMC?

All of these sampling techniques, require **time consuming** evaluations of the a posteriori distribution π for each new draw

- Importance sampling: allow for parallel computation.
- MCMC: can not be parallelized. well, say, most of them

The efficiency of these sampling techniques depend on **design parameters**

- Importance sampling: the proposal distribution.
- Hastings-Metropolis type MCMC: the proposal distribution.

↔ towards **adaptive** algorithms that learn on the fly how to modify the value of the design parameters.

Monitoring convergence

- Importance sampling: criteria such as Effective Sample Size (ESS) or the Normalized Perplexity.
- MCMC: acceptance probability (Hastings-Metropolis algorithms)

Therefore, we decided to

run an **adaptive Importance Sampling** algorithm: **Population Monte Carlo** [Robert et al. 2005]

compare it to an **adaptive MCMC** algorithm: **Adaptive Metropolis algorithm** [Haario et al. 1999]

Population Monte Carlo (PMC) algorithm

- Idea: choose the **best** proposal distribution among a set of (parametric) distributions.
Criterion based on the Kullback-Leibler divergence

$$q_{\star} = \operatorname{argmax}_{q \in \mathcal{Q}} \int \log q(x) \pi(x) dx$$

- In order to have a / to approximate the solution of this optimization problem
 - choose \mathcal{Q} as the set of mixtures of Gaussian distributions (or t -distributions).
 - solve the optimization problem

LOOK! EM algorithm for fitting mixture models on i.i.d. samples $\{Y_k, k \geq 0\}$

$$\operatorname{argmax}_{q \in \mathcal{Q}} \frac{1}{n} \sum_{k=1}^n \log q(Y_k)$$

PMC (II)

How to solve

$$\operatorname{argmax}_{\theta} \int \log \left(\sum_{d=1}^D \alpha_d \mathcal{N}(\mu_d, \Sigma_d)(x) \right) \pi(x) dx \quad \theta = (\alpha_d, \mu_d, \Sigma_d)_{d \leq D}$$

Tool: EM algorithm for mixture models. Given the current estimate $\theta^{(t)}$, update the parameter by

$$\alpha_d^{(t+1)} = \int \rho_d(x; \theta^{(t)}) \pi(x) dx$$

$$\mu_d^{(t+1)} = \frac{1}{\alpha_d^{(t+1)}} \int x \rho_d(x; \theta^{(t)}) \pi(x) dx$$

$$\Sigma_d^{(t+1)} = \frac{1}{\alpha_d^{(t+1)}} \int (x - \mu_d^{(t+1)})(x - \mu_d^{(t+1)})^T \rho_d(x; \theta^{(t)}) \pi(x) dx$$

where

$$\rho_d(x; \theta) = \frac{\alpha_d \mathcal{N}(\mu_d, \Sigma_d)(x)}{\sum_{j=1}^D \alpha_j \mathcal{N}(\mu_j, \Sigma_j)(x)} = \text{prob. of the component } d \text{ given } x$$

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How to solve

$$\operatorname{argmax}_{\theta} \int \log \left(\sum_{d=1}^D \alpha_d \mathcal{N}(\mu_d, \Sigma_d)(x) \right) \pi(x) dx \quad \theta = (\alpha_d, \mu_d, \Sigma_d)_{d \leq D}$$

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$$\operatorname{argmax}_{\theta} \int \log \left(\sum_{d=1}^D \alpha_d \mathcal{N}(\mu_d, \Sigma_d)(x) \right) \pi(x) dx \quad \theta = (\alpha_d, \mu_d, \Sigma_d)_{d \leq D}$$

Tool: EM algorithm for mixture models. Given the current estimate $\theta^{(t)}$, update the parameter by

$$\alpha_d^{(t+1)} = \sum_{k=1}^N \bar{\omega}_k \rho_d(X_k; \theta^{(t)})$$

$$\mu_d^{(t+1)} = \frac{1}{\alpha_d^{(t+1)}} \sum_{k=1}^N \bar{\omega}_k X_k \rho_d(X_k; \theta^{(t)})$$

$$\Sigma_d^{(t+1)} = \frac{1}{\alpha_d^{(t+1)}} \sum_{k=1}^N \bar{\omega}_k (X_k - \mu_d^{(t+1)})(X_k - \mu_d^{(t+1)})^T \rho_d(x; \theta^{(t)})$$

where $\{(\bar{\omega}_k, X_k)\}_k$ is a (normalized) particle approximation of π

PMC (III)

Iterative algorithm:

- **initialization**: choose an initial proposal distribution $q^{(0)}$ and draw weighted points $\{(w_k, X_k)\}_k$ that approximate π
- **Iteration 1**: Based on these samples,
 - update the proposal distribution

$$q^{(1)}(x) = \sum_{d=1}^D \alpha_d^{(1)} \mathcal{N}(\mu_d^{(1)}, \Sigma_d^{(1)})(x)$$

by applying the EM update formula.

- Draw weighted points $\{(w_k, X_k)\}_k$ that approximate π , by importance sampling with proposal $q^{(1)}$.
- **Repeat** until \dots further adaptations do not result in significant improvements of the KL divergence.

PMC - stopping rules (IV)

From the particle approximation $\{(\omega_k, X_k), k \leq N\}$,

- 1 compute the **Normalized Effective Sample Size** at each iteration

$$\text{ESS} = \frac{1}{N} \left(\sum_{k=1}^N \bar{\omega}_k^2 \right)^{-1} \quad \text{where } \bar{\omega}_k = \frac{\omega_k}{\sum_{j=1}^N \omega_j}$$

that can be interpreted as the proportion of sample points with non-zero weights.

- 2 compute the **normalized perplexity**

$$\frac{1}{N} \exp \left(- \sum_{k=1}^N \bar{\omega}_k \log(\bar{\omega}_k) \right)$$

In both cases, values close to 1 indicate good agreement.

Adaptive Metropolis

- Symmetric Random Walk Metropolis algorithm
- with Gaussian proposal distribution, with "mysterious" (but famous) scaling matrix

$$\mathcal{N}\left(0, \frac{2.38^2}{d} \Sigma_\pi\right)$$

where Σ_π is the **unknown** covariance matrix of π . [Roberts et al. 1997]

- "unknown"?! estimate it on the fly, from the samples of the algorithm → adaptive Metropolis algorithm

III. Simulations

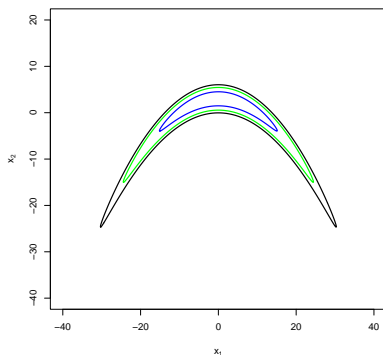
Simulations

on

- 1 simulated data, from a "banana" density
- 2 real data.

Simulated data

The target distribution in \mathbb{R}^{10} . Below marginal distribution of (x_1, x_2)



and (x_3, \dots, x_{10}) are independent $\mathcal{N}(0,1)$.

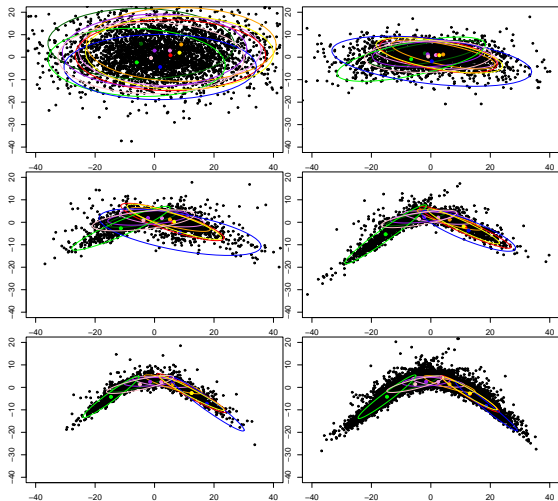


FIG.: Iterations 1,3,5,7,9,11. 10k points per plot, except 100k in the last one. Mixture of 9 t -distributions, with 9 degrees of freedom

Monitoring convergence: the *Normalized perplexity* (top panel) and the *Normalized Effective Sample size* (bottom panel)

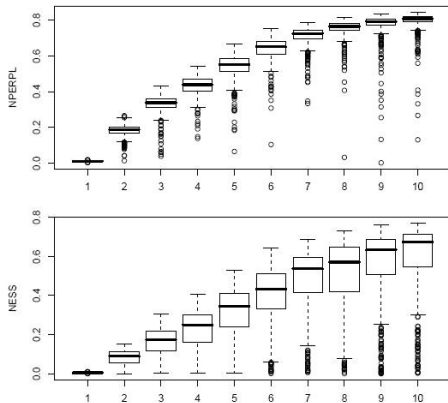


FIG.: for the first 10 iterations, over 500 simulation runs.

Comparison of adaptive MCMC and PMC:

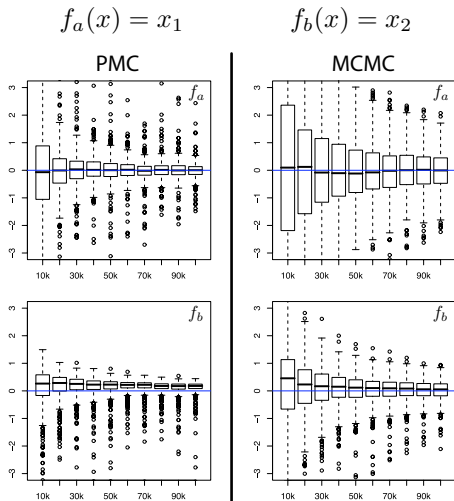


FIG.: for the first 10 iterations, over 500 simulation runs.

Application to cosmology

Evolution of the PMC algorithm: *the likelihood is from the SNIa data*

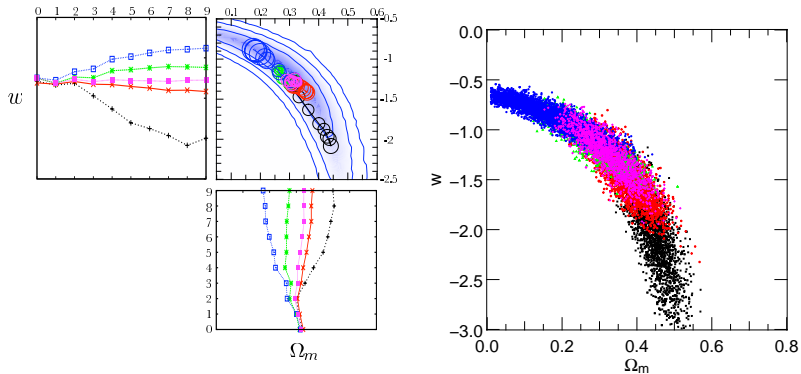


FIG.: [left] evolution of the Gaussian mixtures with 5 components. [right] samples at the last PMC iteration, from the 5 components

Evolution of the weights: *the likelihood is WMAP5 for a flat Λ CDM model with six parameters*

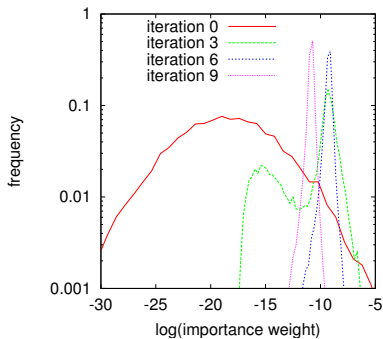


FIG.: Histogram of the normalized weights for four iterations

- Monitoring convergence: *the likelihood is WMAP5 for a flat Λ CDM model with six parameters*

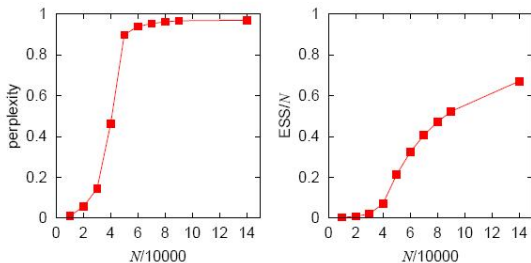


FIG.: perplexity (left) and ESS (right) as a function of the cumulative sample size

- After $150k$ evaluations of π : ESS is about 0.7; mean acceptance rate in MCMC about 0.25.

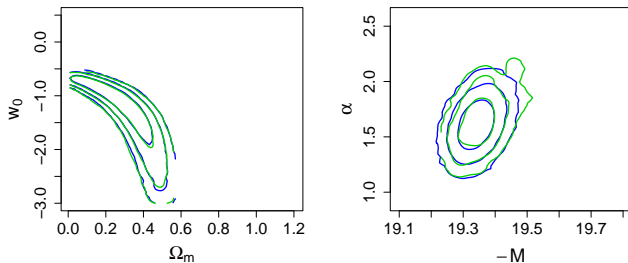
Comparison of MCMC and PMC: *the likelihood is from the SNIa data*

FIG.: Marginalized likelihoods (68%, 95%, 99.7% contours are shown) for PMC (solid blue) and MCMC (dashed green)

Estimates of cosmological parameters: *from the WMAP5 data (left) and from the lensing+SNIa+CMB data sets (right)*

Parameter	PMC	MCMC
Ω_b	$0.04424^{+0.00321}_{-0.00290}$	$0.04418^{+0.00321}_{-0.00294}$
Ω_m	$0.2633^{+0.0340}_{-0.0282}$	$0.2626^{+0.0359}_{-0.0280}$
τ	$0.0878^{+0.0181}_{-0.0160}$	$0.0885^{+0.0181}_{-0.0160}$
n_s	$0.9622^{+0.0145}_{-0.0143}$	$0.9628^{+0.0139}_{-0.0145}$
$10^9 \Delta_R^2$	$2.431^{+0.118}_{-0.113}$	$2.429^{+0.123}_{-0.108}$
h	$0.7116^{+0.0271}_{-0.0261}$	$0.7125^{+0.0274}_{-0.0268}$

Parameter	PMC	MCMC
Ω_b	$0.0432^{+0.0027}_{-0.0024}$	$0.0432^{+0.0026}_{-0.0023}$
Ω_m	$0.254^{+0.018}_{-0.017}$	$0.253^{+0.018}_{-0.016}$
τ	$0.088^{+0.018}_{-0.016}$	$0.088^{+0.019}_{-0.015}$
w	-1.011 ± 0.060	$-1.010^{+0.059}_{-0.060}$
n_s	$0.963^{+0.015}_{-0.014}$	$0.963^{+0.015}_{-0.014}$
$10^9 \Delta_R^2$	$2.413^{+0.098}_{-0.093}$	$2.414^{+0.098}_{-0.092}$
h	$0.720^{+0.022}_{-0.021}$	$0.720^{+0.023}_{-0.021}$
a	$0.648^{+0.040}_{-0.041}$	$0.649^{+0.043}_{-0.042}$
b	$9.3^{+1.4}_{-0.9}$	$9.3^{+1.7}_{-0.9}$
c	$0.639^{+0.084}_{-0.070}$	$0.639^{+0.082}_{-0.070}$
$-M$	19.331 ± 0.030	$19.332^{+0.029}_{-0.031}$
α	$1.61^{+0.15}_{-0.14}$	$1.62^{+0.16}_{-0.14}$
$-\beta$	$-1.82^{+0.17}_{-0.16}$	-1.82 ± 0.16
σ_8	$0.795^{+0.028}_{-0.030}$	$0.795^{+0.030}_{-0.027}$

FIG.: Means and 68% confidence intervals

Conclusion

Cosmology provides challenging problems for Bayesian inference:

- large dimension of the parameter space
- time consuming likelihood

Open questions:

- parallelization of Monte Carlo methods
- methods robust to the dimension

Public release of the Bayesian sampling algorithm for cosmology, CosmoPMC (Martin KILBINGER and Karim BENABED)



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CosmoPMC: Cosmology Population Monte Carlo

Martin Kilbinger, Karim Benabed, Olivier Cappe, Jean-Francois Cardoso, Gersende Fort, Simon Prunet, Christian P. Robert, Darren Wraith

(Submitted on 5 Jan 2011)

We present the public release of the Bayesian sampling algorithm for cosmology, CosmoPMC (Cosmology Population Monte Carlo). CosmoPMC explores the parameter space of various cosmological probes, and also provides a robust estimate of the Bayesian evidence. CosmoPMC is based on an adaptive importance sampling method called Population Monte Carlo (PMC). Various cosmology likelihood modules are implemented, and new modules can be added easily. The importance-sampling algorithm is written in C, and fully parallelised using the Message Passing Interface (MPI). Due to very little overhead, the wall-clock time required for sampling scales approximately with the number of CPUs. The CosmoPMC package contains post-processing and plotting programs, and in addition a Monte-Carlo Markov chain (MCMC) algorithm. The sampling engine is implemented in the library pmclib, and can be used independently. The software is available for download at [this http URL](#).

Comments: CosmoPMC user's guide, version v1.0

Subjects: **Cosmology and Extragalactic Astrophysics (astro-ph.CO)**

Cite as: **arXiv:1101.0950v1 [astro-ph.CO]**

Submission history

From: Martin Kilbinger [[view email](#)]

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References

- **(paper)** D. Wraith et al. Estimation of cosmological parameters using adaptive importance sampling. Phys. Rev. D, 80(2), 2009.
- Data sets
 - (CMB)** J.F. Cardoso. Precision cosmology with the Cosmic Microwave Background. IEEE Signal Processing Magazine, 2010.
 - Data set CMB** G. Hinshaw et al. ApJS 180, 225, 2009.
 - Data set SNIa** P. Astier et al. A&A 447, 31, 2006.
 - Data set cosmic shear** L. Fu. A&A 479, 9, 2008.
- PMC
 - (algo)** O. Cappé et al. Population Monte Carlo. J. Comput. Graph. Statist. 13(4):907-929, 2004.
 - (cvg results)** R. Douc et al. Convergence of adaptive mixtures of importance sampling schemes. Ann. Statist. 35(1):420-448, 2007.