## Convergence and Efficiency of Adaptive Importance Sampling techniques with partial biasing

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Joint work with

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Talk based on the paper

G. Fort, B. Jourdain, T. Lelièvre, G. Stoltz *Convergence and Efficiency of Adaptive Importance Sampling techniques with partial biasing*, arXiv:1610.0919

## Motivation (1/4)

Goal:

Explore the support of a distribution  $\pi d\lambda$  on  $X \subseteq \mathbb{R}^p$ and/or compute integrals w.r.t.  $\pi$ 

$$\int_{\mathsf{X}} f(x) \ \pi(x) \mathsf{d}\lambda(x)$$

when  $\pi$  is highly metastable, p is large.

Solution: based on Importance Sampling (IS)

Sample  $X_1, \cdots, X_n, \cdots \stackrel{i.i.d.}{\sim} \widetilde{\pi} \, \mathrm{d}\lambda$ 

Define the IS approximation

$$\int_{\mathsf{X}} f \, \pi \mathsf{d}\lambda \approx \frac{1}{n} \sum_{k=1}^{n} \underbrace{\frac{\pi(X_k)}{\widetilde{\pi}(X_k)}}_{\text{interpreting optimization optimization}} f(X_k)$$

importance ratio

#### Motivation (2/4) - How to choose $\tilde{\pi}$ ?

• Define a partition of the support X (Molecular dynamics: Chipot, Pohorille (2007) and Lelievre, Rousset, Stoltz

(2010); Statistics: Chopin, Lelievre, Stoltz (2012))

$$\mathsf{X} = igcup_{i=1}^d \mathsf{X}_i \qquad d ext{ strata}$$

• A family of auxiliary distribution based on a local biasing For all *positive vector*  $\tau = (\tau(1), \cdots, \tau(d))$   $\tau_{(i) > 0, \forall i}$ 

$$\pi_{\tau}(x) \stackrel{\text{def}}{=} \frac{1}{\sum_{i=1}^{d} \frac{\theta_{\star}(i)}{\tau(i)}} \sum_{i=1}^{d} \frac{\pi(x)}{\tau(i)} \mathbb{I}_{\mathsf{X}_{i}}(x),$$

where

$$heta_\star(i) \stackrel{
m def}{=} \int_{\mathsf{X}_i} \, \pi \mathsf{d}\lambda, \qquad \mathsf{up} \ \mathsf{to} \ \mathsf{a} \ \mathsf{constant}, \ \log heta_\star(i) \ \mathsf{is} \ \mathsf{the} \ \mathsf{free-energy}$$

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where

$$\theta_{\star}(i) \stackrel{\text{def}}{=} \int_{\mathsf{X}_i} \pi \mathsf{d}\lambda, \qquad \text{up to a constant, } \log \theta_{\star}(i) \text{ is the free-energy}$$

Key property:  $\pi_{\theta_{\star}}(X_i) = 1/d$  – all the strata have the same weight: efficient to tackle multimodality ! but  $\theta_{\star}$  is unknown.

#### Motivation - Adaptive Importance Sampling (3/4)

#### An iterative algorithm which

 $\bullet$  Will learn on the fly the weight vector  $\theta_{\star}$  though a Stochastic Approximation algorithm

$$\theta_{n+1} = \theta_n + \gamma_{n+1} H(\theta_n, X_{n+1})$$

where H is chosen so that  $\theta_{\star}$  is the unique solution of

$$\int H(\theta, x) \ \pi_{\theta}(x) \, \mathsf{d}\lambda(x) = 0.$$

• from draws  $X_{n+1}$ 

$$X_{n+1} \sim P_{\theta_n}(X_n, \cdot)$$
 kernel with inv. dist.  $\pi_{\theta_n}$ 

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$$X_{n+1} \sim P_{\theta_n}(X_n, \cdot)$$
 kernel with inv. dist.  $\pi_{\theta_r}$ 

#### If convergence is established

- An estimator of the free energy:  $\lim_n \theta_n = \theta_{\star}$ .
- An approximatiton of the target distribution  $\pi$  computed on the fly/online

$$\int f \, \pi \mathsf{d}\lambda = \lim_{n} \frac{d}{n} \sum_{k=1}^{n} f(X_k) \left( \sum_{i=1}^{d} \theta_k(i) \mathbb{I}_{\mathsf{X}_i}(X_k) \right)$$

#### Motivation - Choice of the field $H(\theta, x)$ (4/4)

A family of algorithms: Wang Landau, Self Healing Umbrella Sampling (SHUS), Well-Tempered Metadynamics, SHUS $_{\rho}^{g}$ 

on the form

**(**) Given a new draw  $X_{n+1} \sim P_{\theta_n}(X_n, \cdot)$  with inv. dist.  $\pi_{\theta_n}$ 

2 Update a counter of the visits to a stratum

 $C_{n+1}(i) = C_n(i) + (\cdots)^2 \quad \mathbb{I}_{X_i}(X_{n+1}) \qquad i = 1, \cdots, d$ 

Solution Normalize the counter to obtain a weight vector

$$\theta_{n+1}(i) = \frac{C_{n+1}(i)}{\sum_{j=1}^{d} C_{n+1}(j)} = \theta_n(i) + \gamma_{n+1} \dots + O(\gamma_{n+1}^2) \qquad i = 1, \dots, d$$

Fundamental: if  $X_{n+1} \in X_i$ 

$$\begin{split} C_{n+1}(i) > C_n(i), & C_{n+1}(j) = C_n(j), j \neq i \\ \Longrightarrow \pi_{\theta_{n+1}}(\mathsf{X}_i) < \pi_{\theta_n}(\mathsf{X}_i), & \pi_{\theta_{n+1}}(\mathsf{X}_j) = \pi_{\theta_n}(\mathsf{X}_j). \end{split}$$

## A Wang-Landau (WL) based algorithm

#### a WL based algorithm - algorithm (1/3)

(adapted from) the Wang-Landau algorithm (Wang and Landau, 2001) Input:

- initial values: a point  $X_0 \in \mathsf{X}$  and a counter  $C_0 \in (\mathbb{R}^{\star}_+)^d$
- a positive (deterministic) stepsize sequence  $\{\gamma_n, n \ge 0\}$

For  $n = 0, 1, \cdots$ 

- Normalize the counter

$$\theta_n(i) = \frac{C_n(i)}{\sum_{j=1}^d C_n(j)}, \qquad \forall i = 1, \cdots, d$$

- Draw a new point:  $X_{n+1} \sim P_{\theta_n}(X_n, \cdot)$  kernel with inv. dist.  $\pi_{\theta_n}$ 

- Update the counter of the visited stratum

$$C_{n+1}(i) = C_n(i) + \gamma_{n+1} C_n(i) \mathbb{I}_{X_i}(X_{n+1}), \quad \forall i = 1, \cdots, d$$

#### a WL based algorithm - convergence results (2/3)

On the form

$$\theta_{n+1}(i) = \theta_n(i) + \gamma_{n+1} \left( \theta_n(i) \mathbb{1}_{\mathsf{X}_i}(X_{n+1}) - \sum_{j=1}^d \theta_n(j) \mathbb{1}_{\mathsf{X}_j}(X_{n+1}) \right) + \gamma_{n+1}^2 O_{w.p.1.}(1).$$

#### a WL based algorithm - convergence results (2/3)

On the form

$$\theta_{n+1}(i) = \theta_n(i) + \gamma_{n+1} \left( \theta_n(i) \mathbb{1}_{\mathsf{X}_i}(X_{n+1}) - \sum_{j=1}^d \theta_n(j) \mathbb{1}_{\mathsf{X}_j}(X_{n+1}) \right) + \gamma_{n+1}^2 O_{w.p.1.}(1).$$

Under conditions on

- the strata and the target:  $0 < \inf_{\mathsf{X}} \pi \le \sup_{\mathsf{X}} \pi < \infty$ ,  $\theta_{\star}(i) > 0$ .
- the ergodicity of the kernels  $P_{\theta}$
- the stepsize sequence  $\gamma_n$ :  $\sum_n \gamma_n = +\infty$ ,  $\sum_n \gamma_n^2 < \infty$

it is proved asymptotic results (F., Jourdain, Kuhn, Lelièvre, Stoltz, 2015a)

- **1** The a.s. convergence of the sequence  $\theta_n$  to  $\theta_{\star}$ .
- 2 The "convergence" of the samples  $\{X_1, \cdots, X_n, \cdots\}$

$$\int f \, \pi \mathsf{d}\lambda = \lim_{n} \frac{d}{n} \sum_{k=1}^{n} f(X_k) \left( \sum_{i=1}^{d} \frac{\theta_k(i)}{\mathbb{I}_{\mathsf{X}_i}(X_k)} \right) \qquad a.s.$$

 $\hookrightarrow$  very bad Effective Sample Size

#### a WL based algorithm - convergence results (3/3)

and role of the stepsize sequence (F., Jourdain, Kuhn, Lelièvre, Stoltz, 2015b) in the transient phase



Figure: Left: level curves of the target density. Right: typical trajectory for  $\beta = 15$  when  $\gamma_n = \gamma_\star/n^{0.6}$  with  $\alpha = 0.6$  and  $\gamma_\star = 1$ .

• The density depends on a parameter  $\beta$ : large values of  $\beta$  increases the metastability phenomenon.

• We choose 
$$\gamma_n = \gamma_\star/n^\alpha$$
  $\alpha \in (1/2, 1]$   
 $\ln T_{(\alpha < 1)} = C(\alpha, \gamma_\star) + \frac{1}{1 - \alpha} \ln \beta$   $\ln T_{(\alpha = 1)} = C(\gamma_\star) + \frac{\mu_0}{1 + \gamma_\star}\beta$   
• "self tuned" step size  $\gamma_n$ 

## An Adaptive Importance Sampling with

- self-tuned stepsize sequence
- partial biasing to improve the IS step

 $\mathsf{SHUS}^g_\rho$ 

#### A new algorithm

Self-tuned and Partially biasing algorithm (F., Jourdain, Leliévre, Stoltz (2016)) Input:

- initial values: a point  $X_0 \in \mathsf{X}$  and a counter  $C_0 \in (\mathbb{R}^{\star}_+)^d$
- a biasing function ho and a stepsize control function g

For  $n = 0, 1, \cdots$ 

- Normalize the counter

$$\theta_n(i) = \frac{C_n(i)}{\sum_{j=1}^d C_n(j)}, \quad \forall i = 1, \cdots, d$$

- Draw a new point:  $X_{n+1} \sim P_{\rho(\theta_n)}(X_n, \cdot)$ 

- Update the counter of the visited stratum

kernel with inv. dist.  $\pi_{\rho(\theta_n)}$  $\forall i = 1, \cdots, d$ 

$$C_{n+1}(i) = C_n(i) + \frac{\gamma}{g\left(\sum_{j=1}^d C_n(j)\right)} \left(\sum_{j=1}^d C_n(j)\right) \rho\left(\theta_n(i)\right) \mathbb{I}_{\mathsf{X}_i}(X_{n+1}),$$

The samples  $X_n \overset{i.i.d.}{\sim} \pi$ ;

► A counter of the visits to each stratum

$$C_{n}(i) = C_{n-1}(i) + \gamma \mathbb{I}_{X_{i}}(X_{n}) = C_{0}(i) + \gamma \sum_{k=1}^{n} \mathbb{I}_{X_{i}}(X_{k}) \implies C_{n}(i) \sim \gamma n \,\theta_{\star}(i)$$
$$= C_{n-1}(i) + \underbrace{\frac{\gamma}{\sum_{j=1}^{d} C_{n-1}(j)}}_{\gamma_{n} = O(1/n)} \left( \sum_{j=1}^{d} C_{n-1}(j) \right) \quad \mathbb{I}_{X_{i}}(X_{n})$$

 $\blacktriangleright$  The estimate of  $\theta_{\star}$ 

$$\theta_n(i) = \theta_{n-1}(i) + \gamma_n \left( \mathbb{I}_{\mathsf{X}_i}(X_n) - \sum_{j=1}^d \mathbb{I}_{\mathsf{X}_j}(X_n) \right) + O(\gamma_n^2)$$

► For approximation of integrals

$$\int f \pi \mathrm{d}\lambda \approx \frac{1}{n} \sum_{k=1}^{n} f(X_k)$$

The samples  $X_n \overset{i.i.d.}{\sim} \pi$ ;  $X_n \overset{i.i.d.}{\sim} \pi_{\rho(\theta_\star)} \propto \sum_{i=1}^d \frac{\pi}{\rho(\theta_\star(i))} \mathbb{I}_{X_i}$ ;

A counter of the visits to each stratum

$$C_{n}(i) = C_{n-1}(i) + \underbrace{\frac{\gamma}{\sum_{j=1}^{d} C_{n-1}(j)}}_{\gamma_{n} = O(1/n)} \left( \sum_{j=1}^{d} C_{n-1}(j) \right) \rho(\theta_{\star}(i)) \mathbb{I}_{X_{i}}(X_{n})$$

▶ The estimate of  $\theta_{\star}$ 

$$\theta_n(i) = \theta_{n-1}(i) + \gamma_n \left( \rho(\theta_\star(i)) \mathbb{I}_{\mathsf{X}_i}(X_n) - \sum_{j=1}^d \rho(\theta_\star(j)) \mathbb{I}_{\mathsf{X}_j}(X_n) \right) + O_{w.p.1}(\gamma_n^2)$$

► For approximation of integrals

$$\int f \pi \mathsf{d}\lambda \approx \frac{1}{n} \sum_{k=1}^n f(X_k) \left( \sum_{j=1}^d \rho(\theta_\star(j)) \mathbb{I}_{\mathsf{X}_j}(X_k) \right) \left( \sum_{j=1}^d \frac{\theta_\star(j)}{\rho(\theta_\star(j))} \right)$$

The discrepancy between the weights is modified through  $\rho$ . ex.  $t^{a}$ , 0 < a < 1

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► A counter of the visits to each stratum

$$C_n(i) = C_{n-1}(i) + \underbrace{\frac{\gamma}{g\left(\sum_{j=1}^d C_{n-1}(j)\right)}}_{\gamma_n \to 0} \left(\sum_{j=1}^d C_{n-1}(j)\right) \rho(\theta_\star(i)) \ \mathrm{I}_{\mathbf{X}_i}(X_n)$$

▶ The estimate of  $\theta_{\star}$ 

•

$$\theta_n(i) = \theta_{n-1}(i) + \gamma_n \left( \rho(\theta_\star(i)) \mathbb{I}_{\mathsf{X}_i}(X_n) - \sum_{j=1}^d \rho(\theta_\star(j)) \mathbb{I}_{\mathsf{X}_j}(X_n) \right) + O_{w.p.1}(\gamma_n^2)$$

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$$\int f \pi \mathsf{d}\lambda \approx \frac{1}{n} \sum_{k=1}^{n} f(X_k) \left( \sum_{j=1}^{d} \rho(\theta_\star(j)) \mathbb{I}_{\mathsf{X}_j}(X_k) \right) \left( \sum_{j=1}^{d} \frac{\theta_\star(j)}{\rho(\theta_\star(j))} \right)$$

The discrepancy between the weights is modified through  $\rho$ . ex.  $t^{a}$ , 0 < a < 1Control the step size through a function g

The samples  $X_n \overset{i.i.d.}{\sim} \pi$ ;  $X_n \overset{i.i.d.}{\sim} \pi_{\rho(\theta_\star)} \propto \sum_{i=1}^d \frac{\pi}{\rho(\theta_\star(i))} \mathbb{I}_{X_i}$ ; The weight  $\theta_\star$  is learnt along iterations

► A counter of the visits to each stratum

$$C_n(i) = C_{n-1}(i) + \underbrace{\frac{\gamma}{g\left(\sum_{j=1}^d C_{n-1}(j)\right)}}_{\gamma_n \to 0} \left(\sum_{j=1}^d C_{n-1}(j)\right) \rho(\theta_{n-1}(i)) \mathbb{I}_{\mathsf{X}_i}(X_n)$$

▶ The estimate of  $\theta_{\star}$ 

$$\theta_n(i) = \theta_{n-1}(i) + \gamma_n \left( \rho(\theta_{n-1}(i)) \mathbb{1}_{X_i}(X_n) - \sum_{j=1}^d \rho(\theta_{n-1}(j)) \mathbb{1}_{X_j}(X_n) \right) + O_{w.p.1}(\gamma_n^2)$$

► For approximation of integrals

$$\int f \pi \mathrm{d}\lambda \approx \frac{1}{n} \sum_{k=1}^{n} f(X_k) \left( \sum_{j=1}^{d} \rho(\theta_{k-1}(j)) \mathbb{I}_{\mathsf{X}_j}(X_k) \right) \left( \sum_{j=1}^{d} \frac{\theta_{k-1}(j)}{\rho(\theta_{k-1}(j))} \right)$$

The discrepancy between the weights is modified through  $\rho_{\cdot \mbox{ ex }t^a,\,0\,<\,a\,<\,1}$  Control the step size through a function g

#### Assumptions

- **()** On the target density  $0 < \inf_{X} \pi \le \sup_{X} \pi < \infty$  and  $\theta_{\star}(i) > 0$
- ② On the ergodic behavior of the kernels Hastings-Metropolis kernel, with proposal  $q(x, y)d\lambda(y)$  such that  $\inf_{X^2} q > 0$
- $\textbf{O} \text{ On the function } \rho \longrightarrow \text{satisfied with } \rho(t) = t^a \text{ with } a \in [0,1)$
- 0 On the function g, chosen of the form  $g(s)=(\ln(1+s))^{\alpha/(1-\alpha)}$  with  $\alpha\in(1/2,1)$

### Convergence results (1/2)

By using sufficient conditions for convergence of Adaptive MCMC samplers F., Moulines, Priouret (2012) and convergence of Stochastic Approximation algo with controlled Markovian dynamics Andrieu, Moulines, Priouret (2005) Solution of the random sequence  $\gamma_n$  almost-surely,

$$\lim_{n} \gamma_n n^{\alpha} = (1 - \alpha)^{\alpha} \gamma^{1 - \alpha} \left( \sum_{j=1}^{d} \frac{\theta_{\star}(j)}{\rho(\theta_{\star}(j))} \right) \quad \text{a.s.}$$

▶ On the weight sequence  $\theta_n$  almost-surely,

$$\lim_{n} \theta_n = \theta_{\star}$$

▶ On the Importance Sampling step almost-surely,

$$\lim_{n} \frac{1}{n} \sum_{k=1}^{n} f(X_k) \left( \sum_{j=1}^{d} \rho(\theta_{k-1}(j)) \mathbb{1}_{\mathsf{X}_j}(X_k) \right) \left( \sum_{j=1}^{d} \frac{\theta_{k-1}(j)}{\rho(\theta_{k-1}(j))} \right) = \int f \ \pi \mathsf{d}\lambda$$

## Convergence results (2/2)

We wrote the results in the case

$$\begin{split} \rho(t) &= t^a \text{ with } a \in [0,1) \\ g(s) &= (\ln(1+s))^{\alpha/(1-\alpha)} \text{ with } \alpha \in (1/2,1) \end{split}$$

but our convergence analysis also includes the case

- $\rho(t) = t$  and g(s) = s (F., Jourdain, Lelièvre, Stoltz, 2016) In that case, our algorithm is the Self Healing Umbrella Sampling algorithm (Marsili et al. 2006) "no partial biasing" and "self-tuned stepsize"
- $\rho(t) = t^a, a \in [0, 1)$   $g(s) = s^{1-a}$ In that case, our algorithm is a discrete setting of the Well-Tempered metadynamics algorithm (Barducci, Bussi and Parrinello (2008)) "partial biasing" and "self-tuned stepsize" with a correlated parameter a.

# Is there a gain in such a self-tuned and partially biasing algorithm ?



Make the metastability larger by increasing  $\beta$ .

 $\begin{aligned} & \mathsf{Case}\ \rho(t) = t^a\ \mathsf{for}\ a \in [0,1)\\ g(s) = (\ln(1+s))^{\alpha/(1-\alpha)}\ \mathsf{for}\ \alpha \in (1/2,1) \quad \text{in } \sigma_{m} = \sigma_{wp1}(1/n^{\alpha}) \end{aligned}$ 



Figure: Left: Exit times for  $\alpha = 0.8$ . Right: Exit times for  $\alpha = 0.6$ .

Start from the left mode, measure the exit time T i.e. time to reach  $X_{n,1} > 1$ 

- $T \uparrow$  when  $\beta \uparrow$
- for fixed  $\beta$  and a:  $T \downarrow$  when  $\alpha \downarrow$ .
- for fixed  $\beta$  and  $\alpha$ :  $T \downarrow$  when  $a \uparrow$ .
- Linear fit with a slope indep of a:  $\ln T = c + (1 \alpha)^{-1} \ln \beta$

Comparison to the Well-Tempered Metadynamics  $g(s) = s^{1-a} (a) \gamma_n = O(1/n)$  and  $\rho(t) = t^a$  for  $a \in (0, 1)$ 



Figure: Left: Exit times for various values of a. Right: Associated slopes, fitted by 2.43(1-a).

Exit time T

- Linear fit:  $\ln T = c + 2.43(1-a)\beta$
- For fixed  $\beta$ :  $T \downarrow$  when  $a \uparrow$

#### Normalized Effective Sample Size (EF)

Case  $\gamma_n = O(1/n^{\alpha})$  for  $\alpha \in (1/2, 1)$ ,  $\rho(t) = t^a$  for  $a \in [0, 1)$ 



Figure: Efficiency factors EF(a) for various values of  $\beta$ .

$$EF = \frac{\left(n^{-1} \sum_{k=1}^{n} w(X_k)\right)^2}{\left(n^{-1} \sum_{k=1}^{n} w^2(X_k)\right)} \in [0, 1]$$

By definition, when uniform weights, EF = 1.
For fixed β, EF ↑ when a ↓

#### Conclusion

#### A new algorithm

- which estimates the free energy of  $\pi$  by a Stochastic Approximation algorithm, where the stepsize sequence  $\{\gamma_n, n \ge 0\}$  is tuned on the fly
- which provides an approximation of  $\pi$  by a set of weighted points with a controlled discrepancy of the weights.
- $\bullet$  which requires two design parameters  $(\alpha,a)$  to be fixed by the user
  - $\cdot \ a$  close to 1 in the transient phase, and a close to 0 at convergence.
  - $\cdot \ \alpha$  close to 1/2 in the transient phase.
- far more efficient in the transient phase than Well-Tempered Metadynamics or SHUS or WL.