

Stochastic Perturbations of Proximal-Gradient methods for nonsmooth convex optimization: the price of Markovian perturbations

Gersende Fort

LTCI, CNRS and Telecom ParisTech
Paris, France

Based on joint works with

- Eric Moulines (Ecole Polytechnique, France)
- Yves Atchadé (Univ. Michigan, USA)
- Jean-François Aujol (Univ. Bordeaux, France) and Charles Dossal (Univ. Bordeaux, France)

↔ On Perturbed Proximal-Gradient algorithms (2016-v3, arXiv)

Outline

Application: Penalized Maximum Likelihood inference in latent variable models

Stochastic Gradient methods (case $g = 0$)

Stochastic Proximal Gradient methods

Rates of convergence

High-dimensional logistic regression with random effects

Penalized Maximum Likelihood inference, latent variable model

- N observations : $\mathbf{Y} = (Y_1, \dots, Y_N)$
- A negative normalized log-likelihood of the observations \mathbf{Y} , in a latent variable model

$$\theta \mapsto -\frac{1}{N} \log L(\mathbf{Y}, \theta) \quad L(\mathbf{Y}, \theta) = \int p_\theta(x, \mathbf{Y}) \mu(dx)$$

where $\theta \in \Theta \subset \mathbb{R}^d$.

- A penalty term on the parameter θ : $\theta \mapsto g(\theta)$ for sparsity constraints on θ ; usually non-smooth and convex.

Goal: Computation of

$$\theta \mapsto \operatorname{argmin}_{\theta \in \Theta} \left(-\frac{1}{N} \log L(\mathbf{Y}, \theta) + g(\theta) \right)$$

when the *likelihood* L has no closed form expression, and can not be evaluated.

Latent variable model: example (Generalized Linear Mixed Models)

GLMM

- Y_1, \dots, Y_N : indep. observations from a Generalized Linear Model.
- Linear predictor

$$\eta_i = \underbrace{\sum_{k=1}^p X_{i,k} \beta_k}_{\text{fixed effect}} + \underbrace{\sum_{\ell=1}^q Z_{i,\ell} U_\ell}_{\text{random effect}}$$

where

X, Z : covariate matrices

$\beta \in \mathbb{R}^p$: fixed effect parameter

$U \in \mathbb{R}^q$: **random** effect parameter

Latent variable model: example (Generalized Linear Mixed Models)

GLMM

- Y_1, \dots, Y_N : indep. observations from a Generalized Linear Model.
- Linear predictor

$$\eta_i = \underbrace{\sum_{k=1}^p X_{i,k} \beta_k}_{\text{fixed effect}} + \underbrace{\sum_{\ell=1}^q Z_{i,\ell} \mathbf{U}_\ell}_{\text{random effect}}$$

where

X, Z : covariate matrices

$\beta \in \mathbb{R}^p$: fixed effect parameter

$\mathbf{U} \in \mathbb{R}^q$: **random** effect parameter

Example: logistic regression

- Y_1, \dots, Y_N binary independent observations: Bernoulli r.v. with mean $p_i = \exp(\eta_i) / (1 + \exp(\eta_i))$

$$(Y_1, \dots, Y_N) | \mathbf{U} \equiv \prod_{i=1}^N \frac{\exp(Y_i \eta_i)}{1 + \exp(\eta_i)}$$

- Gaussian random effect: $\mathbf{U} \sim \mathcal{N}_q$.

Gradient of the log-likelihood

$$\log L(\mathbf{Y}, \theta) = \log \int p_\theta(x, \mathbf{Y}) \mu(\mathbf{d}x)$$

Under regularity conditions, $\theta \mapsto \log L(\theta)$ is C^1 and

$$\begin{aligned} \nabla_\theta \log L(\mathbf{Y}, \theta) &= \frac{\int \partial_\theta p_\theta(x, \mathbf{Y}) \mu(\mathbf{d}x)}{\int p_\theta(z, \mathbf{Y}) \mu(\mathbf{d}z)} \\ &= \int \partial_\theta \log p_\theta(x, \mathbf{Y}) \underbrace{\frac{p_\theta(x, \mathbf{Y}) \mu(\mathbf{d}x)}{\int p_\theta(z, \mathbf{Y}) \mu(\mathbf{d}z)}}_{\text{the a posteriori distribution}} \end{aligned}$$

Gradient of the log-likelihood

$$\log L(\mathbf{Y}, \theta) = \log \int p_\theta(x, \mathbf{Y}) \mu(\mathrm{d}x)$$

Under regularity conditions, $\theta \mapsto \log L(\theta)$ is C^1 and

$$\begin{aligned} \nabla_\theta \log L(\mathbf{Y}, \theta) &= \frac{\int \partial_\theta p_\theta(x, \mathbf{Y}) \mu(\mathrm{d}x)}{\int p_\theta(z, \mathbf{Y}) \mu(\mathrm{d}z)} \\ &= \int \partial_\theta \log p_\theta(x, \mathbf{Y}) \underbrace{\frac{p_\theta(x, \mathbf{Y}) \mu(\mathrm{d}x)}{\int p_\theta(z, \mathbf{Y}) \mu(\mathrm{d}z)}}_{\text{the a posteriori distribution}} \end{aligned}$$

The gradient of the log-likelihood

$$\nabla_\theta \left\{ -\frac{1}{N} \log L(\mathbf{Y}, \theta) \right\} = \int H_\theta(x) \pi_\theta(\mathrm{d}x)$$

is an untractable expectation w.r.t. the conditional distribution of the latent variable given the observations \mathbf{Y} . For all (x, θ) , $H_\theta(x)$ can be evaluated.

Approximation of the gradient

$$\nabla_{\theta} \left\{ -\frac{1}{N} \log L(\mathbf{Y}, \theta) \right\} = \int_{\mathcal{X}} H_{\theta}(x) \pi_{\theta}(\mathrm{d}x)$$

- 1 Quadrature techniques: poor behavior w.r.t. the dimension of \mathcal{X}
- 2 Monte Carlo approximation with i.i.d. samples: not possible, in general.
- 3 Markov chain Monte Carlo approximations: sample a Markov chain $\{X_{m,\theta}, m \geq 0\}$ with stationary distribution $\pi_{\theta}(\mathrm{d}x)$ and set

$$\int_{\mathcal{X}} H_{\theta}(x) \pi_{\theta}(\mathrm{d}x) \approx \frac{1}{M} \sum_{m=1}^M H_{\theta}(X_{m,\theta})$$

Approximation of the gradient

$$\nabla_{\theta} \left\{ -\frac{1}{N} \log L(\mathbf{Y}, \theta) \right\} = \int_{\mathcal{X}} H_{\theta}(x) \pi_{\theta}(\mathrm{d}x)$$

- ① Quadrature techniques: poor behavior w.r.t. the dimension of \mathcal{X}
- ② Monte Carlo approximation with i.i.d. samples: not possible, in general.
- ③ Markov chain Monte Carlo approximations: sample a Markov chain $\{X_{m,\theta}, m \geq 0\}$ with stationary distribution $\pi_{\theta}(\mathrm{d}x)$ and set

$$\int_{\mathcal{X}} H_{\theta}(x) \pi_{\theta}(\mathrm{d}x) \approx \frac{1}{M} \sum_{m=1}^M H_{\theta}(X_{m,\theta})$$

Stochastic approximation of the gradient

- *a biased approximation*

$$\mathbb{E} \left[\frac{1}{M} \sum_{m=1}^M H_{\theta}(X_{m,\theta}) \right] \neq \int H_{\theta}(x) \pi_{\theta}(\mathrm{d}x).$$

- *if the chain is ergodic "enough", the bias vanishes when $M \rightarrow \infty$.*

To summarize,

Problem:

$$\operatorname{argmin}_{\theta \in \Theta} F(\theta) \quad \text{with } F(\theta) = f(\theta) + g(\theta)$$

when

- $\theta \in \Theta \subseteq \mathbb{R}^d$
- g convex non-smooth function (explicit).
- f is C^1 and its gradient is of the form

$$\nabla f(\theta) = \int H_\theta(x) \pi_\theta(dx) \approx \frac{1}{M} \sum_{m=1}^M H_\theta(X_{m,\theta})$$

where $\{X_{m,\theta}, m \geq 0\}$ is the output of a MCMC sampler with target π_θ .

To summarize,

Problem:

$$\operatorname{argmin}_{\theta \in \Theta} F(\theta) \quad \text{with } F(\theta) = f(\theta) + g(\theta)$$

when

- $\theta \in \Theta \subseteq \mathbb{R}^d$
- g convex non-smooth function (explicit).
- f is C^1 and its gradient is of the form

$$\nabla f(\theta) = \int H_\theta(x) \pi_\theta(dx) \approx \frac{1}{M} \sum_{m=1}^M H_\theta(X_{m,\theta})$$

where $\{X_{m,\theta}, m \geq 0\}$ is the output of a MCMC sampler with target π_θ .

Difficulties:

- **biased** stochastic perturbation of the gradient
- gradient-based methods in the Stochastic Approximation framework (a **fixed** number of Monte Carlo samples)
- weaker conditions on the stochastic perturbation.

Outline

Application: Penalized Maximum Likelihood inference in latent variable models

Stochastic Gradient methods (case $g = 0$)

Stochastic Proximal Gradient methods

Rates of convergence

High-dimensional logistic regression with random effects

Perturbed gradient algorithm

Algorithm:

Given a stepsize/learning rate sequence $\{\gamma_n, n \geq 0\}$:

Initialisation: $\theta_0 \in \Theta$

Repeat:

- compute H_{n+1} , an approximation of $\nabla f(\theta_n)$
- set $\theta_{n+1} = \theta_n - \gamma_{n+1} H_{n+1}$.

M. Benaïm. Dynamics of stochastic approximation algorithms. Séminaire de Probabilités de Strasbourg (1999)

A. Benveniste, M. Métivier and P. Priouret, Adaptive Algorithms and Stochastic Approximations, Springer-Verlag, New York, 1990.

V. Borkar. Stochastic Approximation: a dynamical systems viewpoint. Cambridge Univ. Press (2008).

M. Dufo, Random Iterative Systems, Appl. Math. 34, Springer-Verlag, Berlin, 1997.

H. Kushner, G. Yin. Stochastic Approximation and Recursive Algorithms and Applications. Springer Book (2003).

Sufficient conditions for the convergence

$$\text{Set } \mathcal{L} = \{\theta \in \Theta : \nabla f(\theta) = 0\}, \quad \eta_{n+1} = H_{n+1} - \nabla f(\theta_n).$$

Theorem (Andrieu-Moulines-Priouret(2005); F.-Moulines-Schreck-Vihola(2016))

Assume

- *the level sets of f are compact subsets of Θ and \mathcal{L} is in a level set of f .*
- $\sum_n \gamma_n = +\infty$ and $\sum_n \gamma_n^2 < \infty$.
- $\sum_n \gamma_n \eta_{n+1} \mathbb{1}_{\theta_n \in \mathcal{K}} < \infty$ for any compact subset \mathcal{K} of Θ .

Then

- there exists a compact subset \mathcal{K}_* of Θ s.t. $\theta_n \in \mathcal{K}_*$ for all n .*
- $\{f(\theta_n), n \geq 0\}$ converges to a connected component of $f(\mathcal{L})$.*

If in addition ∇f is locally lipschitz and $\sum_n \gamma_n^2 \|\eta_n\|^2 \mathbb{1}_{\theta_n \in \mathcal{K}} < \infty$, then $\{\theta_n, n \geq 0\}$ converges to a connected component of $\{\theta : \nabla f(\theta) = 0\}$.

When H_{n+1} is a Monte Carlo approximation (1)

$$\nabla f(\theta_n) = \int H_{\theta_n}(x) \pi_{\theta_n}(\mathrm{d}x)$$

Two strategies:

(1) Stochastic Approximation (fixed batch size)

$$H_{n+1} = H_{\theta_n}(X_{1,n}),$$

(2) Monte Carlo assisted optimization (increasing batch size)

$$H_{n+1} = \frac{1}{M_{n+1}} \sum_{m=1}^{M_{n+1}} H_{\theta_n}(X_{m,n}),$$

where $\{X_{m,n}\}_m$ "approximate" the target $\pi_{\theta_n}(\mathrm{d}x)$.

When H_{n+1} is a Monte Carlo approximation (2)

$$\nabla f(\theta_n) = \int H_{\theta_n}(x) \pi_{\theta_n}(dx)$$

- With i.i.d. Monte Carlo:

$$\mathbb{E}[H_{n+1}|\mathcal{F}_n] = \nabla f(\theta_n) \quad \text{unbiased approximation}$$

- With Markov chain Monte Carlo approximation

$$\mathbb{E}[H_{n+1}|\mathcal{F}_n] \neq \nabla f(\theta_n) \quad \text{Biased approximation !}$$

When H_{n+1} is a Monte Carlo approximation (2)

$$\nabla f(\theta_n) = \int H_{\theta_n}(x) \pi_{\theta_n}(dx)$$

- With i.i.d. Monte Carlo:

$$\mathbb{E}[H_{n+1}|\mathcal{F}_n] = \nabla f(\theta_n) \quad \text{unbiased approximation}$$

- With Markov chain Monte Carlo approximation

$$\mathbb{E}[H_{n+1}|\mathcal{F}_n] \neq \nabla f(\theta_n) \quad \text{Biased approximation !}$$

and the bias:

$$|\mathbb{E}[H_{n+1}|\mathcal{F}_n] - \nabla f(\theta_n)| = O_{L^p} \left(\frac{1}{M_{n+1}} \right)$$

does not vanish when the size of the batch is fixed.

When H_{n+1} is a Monte Carlo approximation (3)

$$\theta_{n+1} = \theta_n - \gamma_{n+1} H_{n+1}$$

$$H_{n+1} = \frac{1}{M_{n+1}} \sum_{j=1}^{M_{n+1}} H_{\theta_n}(X_{j,n}) \approx \nabla f(\theta_n)$$

MCMC approx. and fixed batch size

$$\sum_n \gamma_n = +\infty \quad \sum_n \gamma_n^2 < \infty \quad \sum_n |\gamma_{n+1} - \gamma_n| < \infty$$

i.i.d. MC approx. / MCMC approx with increasing batch size

$$\sum_n \gamma_n = +\infty \quad \sum_n \frac{\gamma_n^2}{M_n} < \infty \quad \sum_n \frac{\gamma_n}{M_n} < \infty \text{ (case MCMC)}$$

A remark on the proof

$$\begin{aligned} \sum_{n=1}^N \gamma_{n+1} (H_{n+1} - \nabla f(\theta_n)) &= \sum_{n=1}^N \gamma_{n+1} \left(\underbrace{\Delta_{n+1}}_{\text{martingale increment}} + \underbrace{R_{n+1}}_{\text{remainder term}} \right) \\ &= \text{Martingale} + \text{Remainder} \end{aligned}$$

How to define Δ_{n+1} ?

unbiased MC approx

$$\Delta_{n+1} = H_{n+1} - \nabla f(\theta_n)$$

biased MC approx with **increasing** batch size

$$\Delta_{n+1} = H_{n+1} - \mathbb{E}[H_{n+1} | \mathcal{F}_n]$$

biased MC approx with **fixed** batch size

technical !

Stochastic Approximation with MCMC inputs: see e.g.

Benveniste-Metivier-Priouret (1990) Springer-Verlag.

Duflo (1997) Springer-Verlag.

Andrieu-Moulines-Priouret (2005) SIAM Journal on Control and Optimization.

F.-Moulines-Priouret (2012) Annals of Statistics.

F.-Jourdain-Lelièvre-Stoltz (2015,2016) Mathematics of Computation, Statistics and Computing.

F.-Moulines-Schreck-Vihola (2016) SIAM Journal on Control and Optimization.

Outline

Application: Penalized Maximum Likelihood inference in latent variable models

Stochastic Gradient methods (case $g = 0$)

Stochastic Proximal Gradient methods

Rates of convergence

High-dimensional logistic regression with random effects

Problem:

A gradient-based method for solving

$$\operatorname{argmin}_{\theta \in \Theta} F(\theta) \quad \text{with } F(\theta) = f(\theta) + g(\theta)$$

when

- g is non-smooth and convex
- f is C^1 and

$$\nabla f(\theta) = \int_{\mathbf{x}} H_{\theta}(x) \pi_{\theta}(\mathrm{d}x).$$

- Available: Monte Carlo approximation of $\nabla f(\theta)$ through Markov chain samples.

The setting, hereafter

$$\operatorname{argmin}_{\theta \in \Theta} F(\theta) \quad \text{with } F(\theta) = f(\theta) + g(\theta)$$

where

- the function $g: \mathbb{R}^d \rightarrow [0, \infty]$ is **convex, non smooth**, not identically equal to $+\infty$, and lower semi-continuous
- the function $f: \mathbb{R}^d \rightarrow \mathbb{R}$ is a **smooth convex function**
i.e. f is continuously differentiable and there exists $L > 0$ such that

$$\|\nabla f(\theta) - \nabla f(\theta')\| \leq L \|\theta - \theta'\| \quad \forall \theta, \theta' \in \mathbb{R}^d$$

- $\Theta \subseteq \mathbb{R}^d$ is the domain of g : $\Theta = \{\theta : g(\theta) < \infty\}$.

The proximal-gradient algorithm

The Proximal Gradient algorithm

$$\theta_{n+1} = \text{Prox}_{\gamma_{n+1}, g}(\theta_n - \gamma_{n+1} \nabla f(\theta_n))$$

where

$$\text{Prox}_{\gamma, g}(\tau) = \underset{\theta \in \Theta}{\text{argmin}} \left(g(\theta) + \frac{1}{2\gamma} \|\theta - \tau\|^2 \right)$$

Proximal map: Moreau(1962); Parikh-Boyd(2013);

Proximal Gradient algorithm: Nesterov(2004); Beck-Teboulle(2009)

About the Prox-step:

- when $g = 0$: $\text{Prox}(\tau) = \tau$
- when g is the projection on a compact set: the algorithm is the projected gradient.
- in some cases, Prox is explicit (e.g. elastic net penalty). Otherwise, numerical approximation:

$$\theta_{n+1} = \text{Prox}_{\gamma_{n+1}, g}(\theta_n - \gamma_{n+1} \nabla f(\theta_n)) + \epsilon_{n+1}$$

The perturbed proximal-gradient algorithm

The Perturbed Proximal Gradient algorithm

$$\theta_{n+1} = \text{Prox}_{\gamma_{n+1}, g}(\theta_n - \gamma_{n+1} H_{n+1})$$

where H_{n+1} is an approximation of $\nabla f(\theta_n)$.

There exist results under (some of) the assumptions

$$\inf_n \gamma_n > 0, \quad \sum_n \|H_{n+1} - \nabla f(\theta_n)\| < \infty, \quad \text{i.i.d. Monte Carlo approx}$$

i.e. fixed stepsize, increasing batch size and unverifiable conditions for MCMC sampling

Combettes (2001) Elsevier Science.

Combettes-Wajs (2005) Multiscale Modeling and Simulation.

Combettes-Pesquet (2015, 2016) SIAM J. Optim, arXiv

Lin-Rosasco-Villa-Zhou (2015) arXiv

Rosasco-Villa-Vu (2014, 2015) arXiv

Schmidt-Leroux-Bach (2011) NIPS

Convergence of the perturbed proximal gradient algorithm

$$\theta_{n+1} = \text{Prox}_{\gamma_{n+1},g}(\theta_n - \gamma_{n+1} H_{n+1}) \quad \text{with } H_{n+1} \approx \nabla f(\theta_n)$$

$$\text{Set: } \quad \mathcal{L} = \text{argmin}_{\Theta}(f + g) \quad \eta_{n+1} = H_{n+1} - \nabla f(\theta_n)$$

Theorem (Atchadé, F., Moulines (2015))

Assume

- g convex, lower semi-continuous; f convex, C^1 and its gradient is Lipschitz with constant L ; \mathcal{L} is non empty.
- $\sum_n \gamma_n = +\infty$ and $\gamma_n \in (0, 1/L]$.
- Convergence of the series

$$\sum_n \gamma_{n+1}^2 \|\eta_{n+1}\|^2, \quad \sum_n \gamma_{n+1} \eta_{n+1}, \quad \sum_n \gamma_{n+1} \langle S_n, \eta_{n+1} \rangle$$

where $S_n = \text{Prox}_{\gamma_{n+1},g}(\theta_n - \gamma_{n+1} \nabla f(\theta_n))$.

Then there exists $\theta_\star \in \mathcal{L}$ such that $\lim_n \theta_n = \theta_\star$.

When H_{n+1} is a Monte Carlo approximation

$$\theta_{n+1} = \text{Prox}_{\gamma_{n+1}, g}(\theta_n - \gamma_{n+1} H_{n+1})$$

$$H_{n+1} = \frac{1}{M_{n+1}} \sum_{j=1}^{M_{n+1}} H_{\theta_n}(X_{j,n}) \approx \nabla f(\theta_n)$$

MCMC approx. and fixed batch size

$$\sum_n \gamma_n = +\infty \quad \sum_n \gamma_n^2 < \infty \quad \sum_n |\gamma_{n+1} - \gamma_n| < \infty$$

i.i.d. MC approx. / MCMC approx with increasing batch size

$$\sum_n \gamma_n = +\infty \quad \sum_n \frac{\gamma_n^2}{M_n} < \infty \quad \sum_n \frac{\gamma_n}{M_n} < \infty \text{ (case MCMC)}$$

↔ Same conditions as in the Stochastic Gradient algorithm

Outline

Application: Penalized Maximum Likelihood inference in latent variable models

Stochastic Gradient methods (case $g = 0$)

Stochastic Proximal Gradient methods

Rates of convergence

High-dimensional logistic regression with random effects

Problem:

For non negative weights a_k , find an upper bound of

$$\sum_{k=1}^n \frac{a_k}{\sum_{\ell=1}^n a_\ell} F(\theta_k) - \min F$$

It provides

- an upper bound for the cumulative regret ($a_k = 1$)
- an upper bound for an **averaging strategy** when F is convex since

$$F\left(\sum_{k=1}^n \frac{a_k}{\sum_{\ell=1}^n a_\ell} \theta_k\right) - \min F \leq \sum_{k=1}^n \frac{a_k}{\sum_{\ell=1}^n a_\ell} F(\theta_k) - \min F.$$

A deterministic control

Theorem (Atchadé, F., Moulines (2016))

For any $\theta_\star \in \operatorname{argmin}_\Theta F$,

$$\begin{aligned} \sum_{k=1}^n \frac{a_k}{A_n} F(\theta_k) - \min F &\leq \frac{a_0}{2\gamma_0 A_n} \|\theta_0 - \theta_\star\|^2 \\ &+ \frac{1}{2A_n} \sum_{k=1}^n \left(\frac{a_k}{\gamma_k} - \frac{a_{k-1}}{\gamma_{k-1}} \right) \|\theta_{k-1} - \theta_\star\|^2 \\ &+ \frac{1}{A_n} \sum_{k=1}^n a_k \gamma_k \|\eta_k\|^2 - \frac{1}{A_n} \sum_{k=1}^n a_k \langle \mathbf{S}_{k-1} - \theta_\star, \eta_k \rangle \end{aligned}$$

where

$$A_n = \sum_{\ell=1}^n a_\ell, \quad \eta_k = H_k - \nabla f(\theta_{k-1}), \quad \mathbf{S}_k = \operatorname{Prox}_{\gamma_k, g}(\theta_{k-1} - \gamma_k \nabla f(\theta_{k-1})).$$

When H_{n+1} is a Monte Carlo approximation, bound in L^q

$$\left\| F \left(\frac{1}{n} \sum_{k=1}^n \theta_k \right) - \min F \right\|_{L^q} \leq \left\| \frac{1}{n} \sum_{k=1}^n F(\theta_k) - \min F \right\|_{L^q} \leq u_n$$

$$u_n = O(1/\sqrt{n})$$

with fixed size of the batch and (slowly) decaying stepsize

$$\gamma_n = \frac{\gamma_\star}{n^a}, a \in [1/2, 1] \quad M_n = m_\star.$$

With averaging: optimal rate, even with slowly decaying stepsize $\gamma_n \sim 1/\sqrt{n}$.

$$u_n = O(\ln n/n)$$

with increasing batch size and constant stepsize

$$\gamma_n = \gamma_\star \quad M_n = m_\star n.$$

Rate with $O(n^2)$ Monte Carlo samples !

Acceleration (1)

Let $\{t_n, n \geq 0\}$ be a positive sequence s.t.

$$\gamma_{n+1}t_n(t_n - 1) \leq \gamma_n t_{n-1}^2$$

Nesterov acceleration of the Proximal Gradient algorithm

$$\theta_{n+1} = \text{Prox}_{\gamma_{n+1}, g}(\tau_n - \gamma_{n+1} \nabla f(\tau_n))$$

$$\tau_{n+1} = \theta_{n+1} + \frac{t_n - 1}{t_{n+1}} (\theta_{n+1} - \theta_n)$$

Nesterov (1983); Beck-Teboulle (2009)

AllenZhu-Orecchia (2015); Attouch-Peypouquet(2015); Bubeck-TatLee-Singh(2015); Su-Boyd-Candes(2015)

Proximal-gradient

$$F(\theta_n) - \min F = O\left(\frac{1}{n}\right)$$

Accelerated Proximal-gradient

$$F(\theta_n) - \min F = O\left(\frac{1}{n^2}\right)$$

Acceleration (2) Aujol-Dossal-F.-Moulines, work in progress

Perturbed Nesterov acceleration: some convergence results

Choose γ_n, M_n, t_n s.t.

$$\gamma_n \in (0, 1/L], \quad \lim_n \gamma_n t_n^2 = +\infty, \quad \sum_n \gamma_n t_n (1 + \gamma_n t_n) \frac{1}{M_n} < \infty$$

Then there exists $\theta_\star \in \operatorname{argmin}_\Theta F$ s.t $\lim_n \theta_n = \theta_\star$.

In addition

$$F(\theta_{n+1}) - \min F = O\left(\frac{1}{\gamma_{n+1} t_n^2}\right)$$

Schmidt-Le Roux-Bach (2011); Dossal-Chambolle(2014); Aujol-Dossal(2015)

γ_n	M_n	t_n	rate	NbrMC
γ	n^3	n	n^{-2}	n^4
γ/\sqrt{n}	n^2	n	$n^{-3/2}$	n^3

Table: Control of $F(\theta_n) - \min F$

Outline

Application: Penalized Maximum Likelihood inference in latent variable models

Stochastic Gradient methods (case $g = 0$)

Stochastic Proximal Gradient methods

Rates of convergence

High-dimensional logistic regression with random effects

Logistic regression with random effects

The model

- Given $U \in \mathbb{R}^q$,

$$Y_i \sim \mathcal{B} \left(\frac{\exp(x'_i \beta + \sigma z'_i U)}{1 + \exp(x'_i \beta + \sigma z'_i U)} \right), \quad i = 1, \dots, N.$$

- $U \sim \mathcal{N}_q(0, I)$
- Unknown parameters: $\beta \in \mathbb{R}^p$ and $\sigma^2 > 0$.

Stochastic approximation of the gradient of f

$$\nabla f(\theta) = \int H_\theta(u) \pi_\theta(du)$$

with

$$\pi_\theta(u) \propto \mathcal{N}(0, I)[u] \prod_{i=1}^N \frac{\exp(Y_i(x'_i \beta + \sigma z'_i u))}{1 + \exp(x'_i \beta + \sigma z'_i u)}$$

↔ sampled by MCMC Polson-Scott-Windle (2013)

Numerical illustration

- **The Data set** simulated: $N = 500$ observations, a sparse covariate vector $\beta_{\text{true}} \in \mathbb{R}^{1000}$, $q = 5$ random effects.
- **Penalty term** elastic net on β , and $\sigma > 0$.
- **Comparison of 5 algorithms**

Algo1 fixed batch size: $\gamma_n = 0.01/\sqrt{n}$ $M_n = 275$

Algo2 fixed batch size: $\gamma_n = 0.5/n$ $M_n = 275$

Algo3 increasing batch size: $\gamma_n = 0.005$ $M_n = 200 + n$

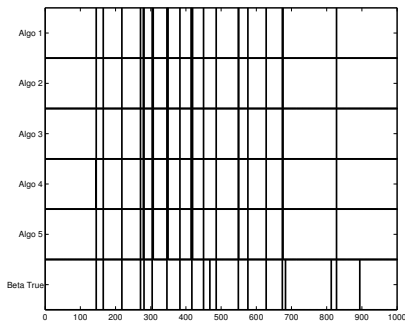
Algo4 increasing batch size: $\gamma_n = 0.001$ $M_n = 200 + n$

Algo5 increasing batch size: $\gamma_n = 0.05/\sqrt{n}$ $M_n = 270 + \sqrt{n}$

After 150 iterations, the algorithms use the same number of MC draws.

A sparse limiting value

Displayed: for each algorithm, the non-zero entries of the limiting value $\beta_\infty \in \mathbb{R}^{1000}$ of a path $(\beta_n)_n$



$$\text{Algo1 } \gamma_n = 0.01/\sqrt{n} \quad M_n = 275$$

$$\text{Algo2 } \gamma_n = 0.5/n \quad M_n = 275$$

$$\text{Algo3 } \gamma_n = 0.005 \quad M_n = 200 + n$$

$$\text{Algo4 } \gamma_n = 0.001 \quad M_n = 200 + n$$

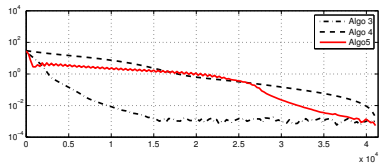
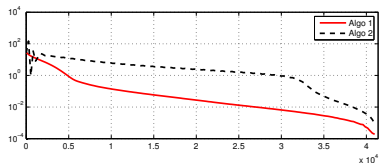
$$\text{Algo5 } \gamma_n = 0.05/\sqrt{n} \quad M_n = 270 + \sqrt{n}$$

Relative error

Displayed: For each algorithm, relative error

$$\frac{\|\beta_n - \beta_{150}\|}{\|\beta_{150}\|}$$

as a function of the total number of MC draws up to time n .



$$(*) \text{ Algo1 } \gamma_n = 0.01/\sqrt{n} \quad M_n = 275$$

$$\text{Algo2 } \gamma_n = 0.5/n \quad M_n = 275$$

$$(*) \text{ Algo3 } \gamma_n = 0.005 \quad M_n = 200 + n$$

$$\text{Algo4 } \gamma_n = 0.001 \quad M_n = 200 + n$$

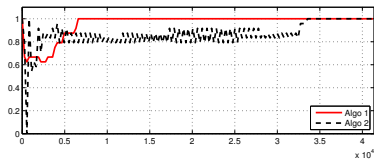
$$\text{Algo5 } \gamma_n = 0.05/\sqrt{n} \quad M_n = 270 + \sqrt{n}$$

Recovery of the sparsity structure of $\beta_\infty (= \beta_{150})$ (1)

Displayed: For each algorithm, the sensitivity

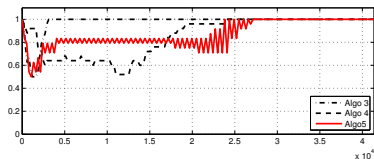
$$\frac{\sum_{i=1}^{1000} \mathbb{I}_{|\beta_{n,i}|>0} \mathbb{I}_{|\beta_{\infty,i}|>0}}{\sum_{i=1}^{1000} \mathbb{I}_{|\beta_{\infty,i}|>0}}$$

as a function of the total number of MC draws up to time n .



$$\begin{aligned} (*) \text{ Algo1 } & \gamma_n = 0.01/\sqrt{n} & M_n &= 275 \\ & \text{Algo2 } & \gamma_n = 0.5/n & M_n &= 275 \end{aligned}$$

$$\begin{aligned} (*) \text{ Algo3 } & \gamma_n = 0.005 & M_n &= 200 + n \\ & \text{Algo4 } & \gamma_n = 0.001 & M_n &= 200 + n \end{aligned}$$



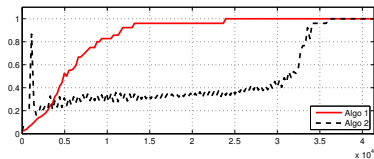
$$\text{Algo5 } \gamma_n = 0.05/\sqrt{n} \quad M_n = 270 + \sqrt{n}$$

Recovery of the sparsity structure of $\beta_\infty (= \beta_{150})$ (2)

Displayed: For each algorithm, the precision

$$\frac{\sum_{i=1}^{1000} \mathbb{I}_{|\beta_{n,i}|>0} \mathbb{I}_{|\beta_{\infty,i}|>0}}{\sum_{i=1}^{1000} \mathbb{I}_{|\beta_{n,i}|>0}}$$

as a function of the total number of MC draws up to time n .

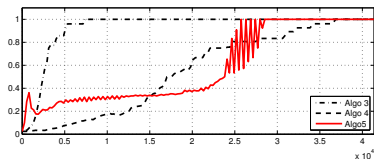


$$(*) \text{ Algo1 } \gamma_n = 0.01/\sqrt{n} \quad M_n = 275$$

$$\text{Algo2 } \gamma_n = 0.5/n \quad M_n = 275$$

$$(*) \text{ Algo3 } \gamma_n = 0.005 \quad M_n = 200 + n$$

$$\text{Algo4 } \gamma_n = 0.001 \quad M_n = 200 + n$$

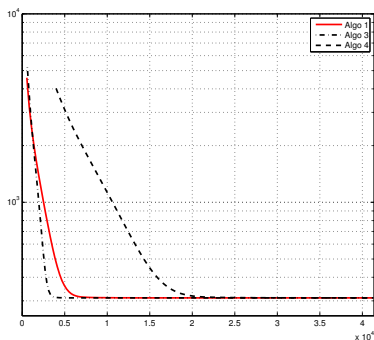


$$\text{Algo5 } \gamma_n = 0.05/\sqrt{n} \quad M_n = 270 + \sqrt{n}$$

Convergence of $\mathbb{E}[F(\theta_n)]$

In this example, the mixed effects are chosen so that $F(\theta)$ can be approximated.

Displayed: For some algorithm, a Monte Carlo approximation of $\mathbb{E}[F(\theta_n)]$ over 50 indep. runs as a function of the total number of MC draws up to time n .



$$(*) \text{ Algo1 } \quad \gamma_n = 0.01/\sqrt{n} \quad M_n = 275$$

$$(*) \text{ Algo3 } \quad \gamma_n = 0.005 \quad M_n = 200 + n$$

$$\text{Algo4 } \quad \gamma_n = 0.001 \quad M_n = 200 + n$$